An optimal control variance reduction method for density estimation

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We study the problem of density estimation of a non-degenerate diffusions using kernel functions. Thanks to Malliavin calculus techniques, we obtain an expansion of the discretisation error. Then, we introduce a new control variate method in order to reduce the variance in the density estimation. We prove a stable law convergence theorem of the type obtained in Jacod-Kurtz-Protter, for the first Malliavin derivative of the error process, which leads us to get a CLT for the new variance reduction algorithm. This CLT gives us a precise description of the optimal parameters of the method.

1 Introduction

In this paper we estimate the density $p(x)$ of a non-degenerate d-dimensional diffusion $(X_t)_{0\leq t\leq T}$ using an Euler scheme X^n of time step T/n. That is, if the diffusion X satisfies the Hörmander condition (see Bally and Talay (1996)) then one obtains the following expansion for the density diffusion

$$p(x) = p_n(x) + \frac{C}{n} + o(1/n),$$

where $p_n(x)$ is a regularized density of the Euler scheme X^n.

In Kohatsu-Higa and Pettersson (2002), a simulation study together with a variance reduction method were introduced. The procedure used can be described as follows.

Consider an integrable continuous function $\phi : \mathbb{R} \to \mathbb{R}$ such that $\int_{\mathbb{R}} \phi(x)dx = 1$ and define the kernels functions

$$\phi_{h,x}(y) = \frac{1}{h}\phi(\frac{y-x}{h}), \quad h > 0 \quad \text{et} \quad x \in \mathbb{R}.$$

Note that $\phi_{h,x} \to \delta_x$ as $h \to 0$, in a weak sense, according to the assumptions on the function ϕ. The idea is then to approximate the density $p(x) = \mathbb{E}\delta_x(X_T)$ by $\mathbb{E}\phi_{h,x}(X_T^n)$ where $h = n^{-\alpha}, \ \alpha > 0$. At this level, a first problem arises. That is, the problem of evaluating the weak error given by

$$\varepsilon_n = \mathbb{E}\phi_{h,x}(X_T^n) - p(x).$$

Kohatsu-Higa and Pettersson (2002) proved that $|\varepsilon_n| \leq C/n$ if $\alpha \geq 1$.

When using this approach a second problem arises, it concerns the problem of the explosion of the variance of the r.v. $\phi_{h,x}(X_T^n)$ when using a Monte Carlo method. In their paper, Kohatsu-Higa and Pettersson (2002) propose then instead the use of the integration by parts formula together with

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a localization method in order to reduce the variance of the method. The asymptotically optimal localization function is found to be of exponential type.

In fact, using the integration by parts formula, Kohatsu-Higa and Pettersson (2002) obtain that

$$\mathbb{E}\phi_{h,x}(X_T^n) = \mathbb{E}\Big(\psi_{h,x}(X_T^n)H_n\Big),$$

where $\psi_{h,x}$ is the primitive function of $\phi_{h,x}$ and H_n is the weight given by the Malliavin calculus. Using this idea, Kohatsu-Higa and Pettersson (2002) construct an efficient control variate which reduces the variance in the Monte carlo estimation of $\mathbb{E}\Big(\psi_{h,x}(X_T^n)H_n\Big)$. The disadvantage of this method is that the computation time of their algorithm is higher than that of classical methods using kernel density functions.

In this work, we propose an alternative approach using the kernel estimation method through the calculation of $\mathbb{E}\phi_{h,x}(X_T^n)$ together with a control method based on the statistical Romberg method (see Kebaier (2005) for more details on the regular case). The method uses two Euler schemes X_T^n and X_T^m with $m << n$ as follows. Simulate a large number, N_m, of sample paths with the coarse time discretization step T/m and few additional sample paths of size N_n with the fine time discretization step T/n.

In this case, in contrast with the regular case studied in Kebaier (2005) there is still explosion of variances. This will be controlled through an appropiate renormalization and a decomposition of the derivatives of the kernel. We will see as a final consequence of Theorem 6.1 that the kernels, as proposed before, in general do not lead to variance reduction. To obtain this variance reduction, one has to consider a subclass of kernel functions known as super kernels of order s where $s > 2(d+1)$ (see Definition 3.1). If fact, otherwise there is no variance reduction with the control method proposed.

As these kernels do not correspond with the original ideas of Bally and Talay (1996), we start by finding the expansion of the weak error ε_n. That is, we prove that

$$\varepsilon_n = \frac{C}{n} + o(1/n)$$

(see theorem 3.1).

As the weak error ε_n is of order $1/n$, we will suppose that all the parameters depends on the time step number n. Hence, we set $h = n^{-\alpha}$, $0 < \alpha < 1/2$ (the window size of the kernel function $\phi_{h,x}$), $m = n^{\beta}$, $0 < \beta < 2/3$ (the time step number of the auxiliary Euler scheme), $N_m = n^{\gamma_1}$, $\gamma_1 > 0$ and $N_n = n^{\gamma_2}$, $\gamma_2 > 0$, where N_m denotes he sample size for the *coarse* estimation of $\mathbb{E}\phi_{h,x}(X_T^n)$ by $\frac{1}{N_m}\sum_{i=1}^{N_m}\phi_{h,x}(X_{T,i}^{n^\beta})$, whereas N_n denotes the sample size needed for the *fine* estimation of $\mathbb{E}\Big\{\phi_{h,x}(X_T^n) - \phi_{h,x}(X_T^{n^\beta})\Big\}$ by $\frac{1}{N_n}\sum_{i=1}^{N_n}\{\phi_{h,x}(X_{T,i}^n) - \phi_{h,x}(X_{T,i}^{n^\beta})\}$.

Our aim is to find the optimal parameters leading to an optimal complexity of the algorithm. In order to obtain these optimal parameters we extend a result of Jacod and Protter (1998) for the asymptotic behavior of the law of the first Malliavin derivative of the error in the Euler scheme. Using this extension we prove a CLT, for our algorithm, giving us a precise description of the choice of the optimal parameters m, N_m and $N_{n,m}$.

The usual version of the integration by parts formula of Malliavin Calculus in dimension d, see Nualart (1995) (p.103, 2006 edition) is based on using d times the integration by parts formula. Although it is feasible to prove the stable convergence of the high order weights, we propose instead to use a new integration by parts formula introduced by Malliavin and Thalmaier (2006) which significantly simplifies the proof in the general multi-dimension context.

The optimal parameters given by the CLT lead to an optimal complexity of the algorithm of order $n^{\frac{5}{2}+(d+1)\alpha}$ which is less than the optimal complexity of the Monte Carlo method which is of order $n^{3+\alpha d}$, where $\alpha \in (0, \frac{1}{2})$ is the parameter tuning the window size h and d is the dimension of the problem. The gain obtained here is of order $n^{\frac{1}{2}-\alpha}$. Consequently, we have an exact mathematical estimate of when and how much variance reduction can be achieved. Whereas, there is less reduction than in the regular case due to the explosion of the variance of our estimators (see section 6 for more details).

The remainder of the paper is organized as follows. In the following section, we introduce some basics of the Malliavin Calculus. In section 3, we study the discretization error ε_n. Section 4 is devoted to prove the CLT for the classical Monte Carlo method. In section 5 we prove a stable convergence theorem for the first Malliavin derivative of the error in the Euler scheme. In the last section we prove a CLT for the statistical Romberg algorithm and we give the optimal parameters leading to an optimal complexity of the method.

In the Appendices we give the proofs of technical lemmas used in the proofs.

2 Malliavin Calculus

2.1 Main definitions and properties

We follow the notations, definitions and results of Nualart (1995). Let $(W_t)_{0\le t\le T}$ be a q dimensional standard Brownian motion defined on the filtered probability space $(\Omega,\mathcal{F},(\mathcal{F}_t),\mathbb{P})$ where $(\mathcal{F}_t)_{0\le t\le T}$ denotes the standard filtration. D denotes the Malliavin derivative which takes values in $H := L^2([0,T];\mathbb{R}^q)$. The k-th order derivative of F for $k\in\{1,...,q\}^{\mathbb{N}}$ is denoted by D^kF taking values in $H^{\otimes |k|}$ and given by

$$D^k_{t_1,\dots,t_{|k|}}F = D^{k_1}_{t_1}\dots D^{k_{|k|}}_{t_{|k|}}F$$

where $|k|$ denotes the length of the multi-index k and k_i, $i=1,...,|k|$ denote its elements.

Note that the operator D^k is closed. For $p\ge 1$ and $k\in\mathbb{N}$, we denote $\mathbb{D}^{k,p}(W)$ the closure of the space of smooth random variables with respect to the norm $\|\cdot\|_{k,p}$.

We denote $\mathbb{D}^\infty(W)=\bigcap_{p\ge1}\bigcap_{k\ge1}\mathbb{D}^{k,p}(W)$. For $F=(F^1,\dots,F^d)\in(\mathbb{D}^\infty(W))^d$, we introduce γ_F the Malliavin covariance matrix of F given by

$$\gamma_F^{ij} = \langle DF^i, DF^j\rangle_H,\quad 1\le i,j\le d$$

2.2 Duality and integration by parts formulas

Let δ denote the adjoint operator of D, which is also called *Skorokhod integral.* The operator δ is closed, we denote by $Dom(\delta)$ its domain (see for example Definition 1.3.1 of Nualart (1995)). Note that if $u\in L^2\big([0,T]\times\Omega;\mathbb{R}^q\big)$ is an adapted process, then (see proposition 1.3.4 Nualart (1995)) $u\in Dom(\delta)$ and $\delta(u)$ coincides with the Itô integral.

If $F\in\mathbb{D}^{1,2}$ and $u\in Dom(\delta)$ then $Fu\in Dom(\delta)$ and we have

$$\delta(Fu) = F\delta(u) - \langle DF, u\rangle_H.$$

In such a case we have the following duality formula

$$\mathbb{E}\Big[\langle u, DF\rangle_H\Big] = \mathbb{E}\big[F\delta(u)\big]. \tag{1}$$

In the following we give the definition of a non-degenerate random vector.

Definition 2.1. *A random vector $F=(F^1,\dots,F^d)\in(\mathbb{D}^\infty(W))^d$ is said to be non-degenerate if the Malliavin covariance matrix of F is invertible a.s. and*

$$(\det\gamma_F)^{-1}\in\bigcap_{p\ge1}L^p(\mathbb{P}^W).$$

For a nondegenerate random vector, the following integration by parts formula plays a key role. (For a proof of the following proposition see Nualart (1998)).

Proposition 2.1. *Let $F\in\big(\mathbb{D}^\infty(W)\big)^d$ be a non-degenerate random vector. Let $f\in C_p^\infty(\mathbb{R}^n)$, and let $G\in\mathbb{D}^\infty(W)$. Fix $k\ge1$. Then for any multi-index $m=(m_1,\dots,m_k)\in\{1,\dots,d\}^k$ we have*

$$\mathbb{E}\big[\partial_m f(F)G\big] = \mathbb{E}\big[f(F)\,\mathbf{H}_m(F,G)\big],$$

where $\partial_m = \partial_{m_1} \dots \partial_{m_k}$ *and the random variable* $\mathbf{H}_m(F,G)$ *is defined inductively as follows*

$$\begin{aligned}\mathbf{H}_{(i)}(F,G) &= \textstyle\sum_{j=1}^d \delta\Big(DF^j G(\gamma_F^{-1})^{ij}\Big)\\ \mathbf{H}_m(F,G) &= \mathbf{H}_{(m_k)}\big(F, \mathbf{H}_{(m_1,\dots,m_{k-1})}(F,G)\big).\end{aligned}$$

2.3 An extension of the integration by parts formula

In the following work we will deal with a d-dimensional diffusion $X = (X^1,\dots,X^d)$ driven by a q-dimensional Brownian motion $W = (W^1,\dots,W^q)$. In order to regularize the Euler scheme associated to the diffusion X, we will employ d additional noises, corresponding to $X^1,\dots,X^d$. In order to do that, we consider a d-dimensional Brownian motion $\bar{W} = (W^{q+1},\dots,W^{q+d})$, independent of $W = (W^1,\dots,W^q)$, and we set

$$\tilde{W} = (W,\bar{W}) = (W^1,\dots,W^q,W^{q+1},\dots,W^{q+d}).$$

Therefore our random vectors are defined on the Wiener space of dimension $r = q+d$, but we should distinguish between the two Brownian motions W et $\bar{W}$ which play different roles in our calculation: W drive the diffusion whereas $\bar{W}$ is an additional noise used for the regularization. Hence, by using again the notations of the preceding subsection we obtain

$$\tilde{D} = (D,\bar{D}) = (D^1,\dots,D^q,D^{q+1},\dots,D^{q+d})$$

and for $\tilde{u} = (u,\bar{u}) = (u^1,\dots,u^q,u^{q+1},\dots,u^{q+d})$ we have

$$\tilde{\delta}(\tilde{u}) = \delta(u) + \bar{\delta}(\bar{u}).$$

The norms $\|F\|_{k,p}$ are norms defined on $\mathbb{D}^{k,p}(\tilde{W})$, thus it involves all the derivatives $\tilde{D} = (D,\bar{D})$. Similarly, the Malliavin covariance matrix of the random vector F is given by

$$\tilde{\gamma}_F = \langle \tilde{D}F, \tilde{D}F\rangle.$$

The auxiliary noise, that we will use, is given by the random vector

$$Z_{n,\theta} := \frac{\bar{W}_T}{n^{\frac{1}{2}+\theta}}, \quad \theta \geq 0.$$

In the following, we introduce the random vector $F = (F_1,\dots,F_d)$ which depends only on $W = (W^1,\dots,W^q)$ and the random variable G which depends only on $\tilde{W} = (W,\bar{W})$. The proposition below, proved by Kohatsu-Higa and Pettersson (2002), gives us an explicit writing of $\tilde{H}_i$ which appears in the integration by parts formula.

Proposition 2.2. *Let* $F \in \big(\mathbb{D}^\infty(W)\big)^d$ *be a non-degenerate random vector. Let* $f \in C_p^\infty(\mathbb{R}^d)$*, and let* $G \in \mathbb{D}^{1,2}(\tilde{W})$*. Fix* $k \geq 1$*. Then for any multi-index* $m = (m_1,\dots,m_k) \in \{1,\dots,d\}^k$ *we have*

$$\mathbb{E}\big[\partial_m f(F+Z_{n,\theta})G\big] = \mathbb{E}\big[f(F+Z_{n,\theta})\,\tilde{\mathbf{H}}_m(F,G)\big], \tag{2}$$

where the random variable $\tilde{\mathbf{H}}_m(F,G)$ *is given by*

$$\begin{aligned}\tilde{\mathbf{H}}_{(i)}(F,G) &= \sum_{j=1}^d \tilde{\delta}\Big(\tilde{D}(F+Z_{n,\theta})^j G(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{ij}\Big)\\ &= \sum_{j=1}^d \delta\Big(G(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{ij}DF^j\Big) + \frac{1}{n^{\frac{1}{2}+\theta}}\sum_{j=1}^d \bar{\delta}\Big(G(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{ij}\bar{D}\bar{W}_T^j\Big),\\ \tilde{\mathbf{H}}_m(F,G) &= \tilde{\mathbf{H}}_{(m_k)}\big(F, \tilde{\mathbf{H}}_{(m_1,\dots,m_{k-1})}(F,G)\big),\end{aligned}$$

with $\bar{\delta}$ *and* $\tilde{\delta}$ *are respectively the adjoint operators of* $\bar{D}$ *and* $\tilde{D}$*.*

3 Weak convergence of the approximate density

Let $(X_t)_{0\le t\le T}$ be a $\mathbb{R}^d$-valued diffusion process which is the solution of the following stochastic differential equation

$$dX_t = b(X_t)dt + \sigma(X_t)dW_t, \quad X_0 = x \in \mathbb{R}^d, \tag{3}$$

where $W = (W^1,\dots,W^q)$ is a q-dimensional Brownian motion defined on the filtered probability space $\mathcal{B} = (\Omega, \mathcal{F}, (\mathcal{F}_t)_{t\ge 0}, P)$, where $(\mathcal{F}_t)_{t\ge 0}$ denotes a filtration satisfying the usual conditions. The functions $b : \mathbb{R}^d \longrightarrow \mathbb{R}^d$ and $\sigma : \mathbb{R}^d \longrightarrow \mathbb{R}^{d\times q}$ are of class $\mathscr{C}_b^{d+3}$. In what follows we denote for $0 \le k \le q$

$$f(X_t) = \begin{pmatrix} b_1(X_t) & \sigma_{11}(X_t) & \dots & \sigma_{1q}(X_t) \\ b_2(X_t) & \sigma_{21}(X_t) & \dots & \sigma_{2q}(X_t) \\ \vdots & \vdots & & \vdots \\ b_d(X_t) & \sigma_{d1}(X_t) & \dots & \sigma_{dq}(X_t) \end{pmatrix}, \; dY_t := \begin{pmatrix} dt \\ dW_t^1 \\ \vdots \\ dW_t^q \end{pmatrix} \text{ and } Y^k = Y_{k+1}, \;\; f_k := \begin{pmatrix} f_{1,k+1} \\ \vdots \\ f_{d,k+1} \end{pmatrix}.$$

Therefore the stochastic differential equation (3) becomes:

$$dX_t = f(X_t)dY_t.$$

The Euler scheme, denoted by X^n, associated to the diffusion X and with discretization step $\delta = T/n$ is defined as:

$$dX_t^n = f(X^n_{\eta_n(t)})dY_t, \quad \eta_n(t) = [t/\delta]\delta.$$

The next result gives bounds on the error of the Euler scheme in the sense of $\|\;\|_{k,p}$-norms. For a proof of this result see Kusuoka and Stroock (1984) and Hu and Watanabe (1996).

Proposition 3.1. *With the previous notation, the following two properties are valid:*

$\mathbf{P}_1$) $\forall t > 0, \quad X_t^n \in \mathbb{D}^\infty$

$\mathbf{P}_2$) $\forall p > 1,\ \forall k \in \mathbb{N}^*, \exists K > 0$ *such that:*

$$\sup_{t\in[0,T]} \|X_T\|_{k,p} + \sup_{t\in[0,T]} \|X_T^n\|_{k,p} \le K(1 + \|x\|) \tag{4}$$

and

$$\sup_{t\in[0,T]} \|X_T^n - X_T\|_{k,p} \le \frac{K}{\sqrt{n}}. \tag{5}$$

Notation:
For a function $V : \mathbb{R}^d \longrightarrow \mathbb{R}^d$, we denote by $\mathbf{D}V$ the Jacobian matrix of V and by $\mathbf{D^2}V$, its Hessian matrix. We suppose that the d-dimensional diffusion process $(X_t)_{0\le t\le T}$, which is the solution of (3) has coefficients σ and b, which satisfy the Hörmander condition (see Section 2.3.2 of Nualart (1995)).

Therefore X admits a smooth density $p_T(x_0, x)$ (see Kusuoka and Stroock (1985)) and in order to simplify the notation, we denote

$$p_T(x_0, x) := p(x).$$

We consider the continuous Euler scheme X^n, with discretization step $\delta = T/n$, defined by:

$$dX_t^n = b(X_{\eta_n(t)})dt + \sigma(X_{\eta_n(t)})dW_t, \quad \eta_n(t) = [t/\delta]\delta.$$

We note here that the Hörmander condition is not enough to guarantee that the Malliavin covariance matrix associated to the Euler scheme X^n, is invertible (this would be true under an ellipticity condition).

To deal with this problem we will regularize the Euler scheme using $X^n + Z_{n,\theta}$ instead of X^n, $Z_{n,\theta}$ denotes a independent random variable defined in Section 2.3 through the relation

$$Z_{n,\theta} = \frac{\tilde{W}_T}{n^{\frac{1}{2}+\theta}}$$

where $\tilde{W}$ is a d-dimensional Brownian motion independent of W. Then we have the following result.

Proposition 3.2. *For $\lambda \in [0,1]$ we introduce*

$$X_T^{n,\lambda} = X_T + \lambda(X_T^n - X_T).$$

Then for all $p \geq 1$ there exists a constant $K_T > 0$ and parameters $p', p'' \geq 1$ such that

$$\sup_n \Big\| \left(\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}}\right)^{-1} \Big\|_p \leq K_T \Big\| (\det \gamma_{X_T})^{-1} \Big\|_{L^{p'}}^{p''} < \infty.$$

Proof. We have that $\mathbb{E}\left(\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}}\right)^{-p} = A_n + B_n$ with

$$A_n := \mathbb{E}\left\{ \left(\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}}\right)^{-p} \mathbf{1}_{\left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right| < \frac{1}{2}\det \gamma_{X_T}} \right\}$$

and

$$B_n := \mathbb{E}\left\{ \left(\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}}\right)^{-p} \mathbf{1}_{\left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right| \geq \frac{1}{2}\det \gamma_{X_T}} \right\}$$

As the diffusion X is non-degenerated in the sense of definition 2.1, we deduce that

$$\sup_n A_n \leq 2^p \mathbb{E}\left(\det \gamma_{X_T}\right)^{-p} < +\infty.$$

On the other hand, we have that

$$\gamma_{X_T^{n,\lambda}+Z_{n,\theta}} = \gamma_{X_T^{n,\lambda}} + \frac{T}{n^{1+2\theta}} I_d.$$

As $\gamma_{X_T^{n,\lambda}}$ is a positive definite matrix we deduce that

$$\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} \geq \left(\frac{T}{n^{1+2\theta}}\right)^d.$$

Therefore, one obtains that

$$B_n \leq \left(\frac{T}{n^{1+2\theta}}\right)^{-dp} \mathbb{P}\Big(\left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right| \geq \frac{1}{2}\det \gamma_{X_T}\Big).$$

Therefore using the Markov inequality, we have that

$$\begin{aligned} B_n &\leq 2^k \left(\frac{T}{n^{1+2\theta}}\right)^{-dp} \mathbb{E}\Big\{ (\det \gamma_{X_T})^{-1} \left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right| \Big\}^k \\ &\leq 2^k \left(\frac{T}{n^{1+2\theta}}\right)^{-dp} \Big\| \left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right|^k \Big\|_{L^2} \left\| (\det \gamma_{X_T})^{-k} \right\|_{L^2} \end{aligned}$$

Therefore from the inequalities (4) and (5), we obtain that

$$\Big\| \left|\det \gamma_{X_T^{n,\lambda}+Z_{n,\theta}} - \det \gamma_{X_T}\right|^k \Big\|_{L^2} \leq \frac{C_k}{n^{\frac{k}{2}}}$$

where C_k is a given constant. Finally, if we take $k = 2dp(1+2\theta)$ we obtain that

$$\sup_n B_n < \infty.$$

□

In what follows we are interested in considering the approximation of the marginal density $p(x)$ of the diffusion X using kernel density estimation methods.

Definition 3.1. *Let $\phi \in C_b^\infty(\mathbb{R};\mathbb{R})$, we say that ϕ is a super-kernel of order $s > 2$ if*

$$\int_{\mathbb{R}} \phi(x)\,dx = 1, \qquad \int_{\mathbb{R}} x^i \phi(x)\,dx = 0, \quad \forall\, i = 1,\dots s-1, \quad \textit{and} \quad \int_{\mathbb{R}} x^s \phi(x)\,dx \neq 0.$$

In what follows, we suppose that ϕ satisfies the following properties:

a) $\displaystyle\int_{\mathbb{R}} |x|^{s+1} |\phi(x)|\,dx < \infty$, where s denotes the order of the kernel,

b) $\displaystyle\int_{\mathbb{R}} |\phi'(x)|^2\,dx < \infty$.

For $h > 0$, we define

$$\phi_{h,x}(y) = \frac{1}{h}\phi\Big(\frac{y-x}{h}\Big).$$

The parameter h is called *the window size of the kernel.* In the calculations to follows, we will also use other kernels that stem from ϕ. So, we define

$$\phi_{(2),h,x}(y) = \frac{1}{h\phi_{(2)}}\Big[\phi'\Big(\frac{y-x}{h}\Big)\Big]^2, \quad \phi_{i,h,x}(y) = \frac{1}{h\phi_i}\Big|\phi\Big(\frac{y-x}{h}\Big)\Big|^i$$

with

$$\phi_{(2)} = \int |\phi'(x)|^2\,dx \qquad \text{and} \qquad \phi_i := \int |\phi(x)|^i\,dx, \quad \text{for} \quad i = 1,\dots,d.$$

To construct super kernels on $\mathbb{R}^d$, we consider products of unidimensional super kernels.

That is, let $\phi_i : \mathbb{R} \mapsto \mathbb{R}$ for $i = 1,\dots,d$ be given and define

$$\phi(u_1,\dots,u_d) = \phi_1(u_1) \times \cdots \times \phi_d(u_d)$$

and

$$\phi_{h,x}(y) = \frac{1}{h^d}\phi\Big(\frac{y-x}{h}\Big) = \prod_{i=1}^d \phi_{i,h,x}(y_i)$$

We say that ϕ is a super kernel of order s if the functions ϕ_i, $i = 1,\dots,d$ are unidimensional super kernels of order s.

Remark 1. One can construct super kernels of infinite order in the following way. We take a function $\psi \in \mathscr{S}$ (where $\mathscr{S}$ denotes the class of Schwarz tempered distributions) so that $\psi(x) = 1$ in a neighborhood of zero. Next, we define ϕ as the inverse Fourier transform of ψ. That is,

$$\phi(x) := \frac{1}{2\pi}\int_{\mathbb{R}} e^{ix\xi}\psi(\xi)\,d\xi, \quad x \in \mathbb{R}.$$

Then the Fourier transform of ϕ is ψ given by

$$\psi(\xi) = \int_{\mathbb{R}} e^{-ix\xi}\phi(x)\,dx, \quad \xi \in \mathbb{R}.$$

As $\psi^{(k)}(0) = 0$, for all $k \in \mathbb{N}$ we conclude also that $\int_{\mathbb{R}} x^k \phi(x)\,dx = 0$ for all $k \in \mathbb{N}$ and as $\psi(0) = 1$ we have that $\int_{\mathbb{R}} \phi(x)\,dx = 1$. The inverse Fourier transform sends the functions $\mathscr{S}$ into $\mathscr{S}$. Therefore $\phi \in \mathscr{S}$ and consequently, it verifies the conditions a) and b) above.

Also, one can easily contruct polynomials on compacts which lead to super kernels of order s which are not of order $s+1$.

The property that will interest us in the calculations to follow is that the super kernel of order s approximate the Dirac delta function up to the order $s+1$. More precisely, we have the following result.

Lemma 3.1. *1. Let $\phi : \mathbb{R}^d \to \mathbb{R}$ be a super kernel of order s, i.e. a d-dimensional super kernel of the form $\phi(x) = \prod_{j=1}^d \phi_j(x_j)$ where for $j = 1, \dots, d$, $\phi_j : \mathbb{R} \to \mathbb{R}$ denote unidimensional super kernels of order s. Let $f \in \mathcal{C}_b^{s+1}(\mathbb{R}^d; \mathbb{R})$. Then*

$$\Big| f(x) - \int_{\mathbb{R}^d} f(y)\phi_{h,x}(y)\, dy - \frac{h^s}{s!} \sum_{|\alpha|=s} \partial^\alpha f(x) \int_{\mathbb{R}^d} \prod_{i=1}^s u_{\alpha_i} \phi(u)\, du \Big| \le C\, h^{s+1},$$

where $\partial^\alpha f$ denotes the partial derivative of f with respect to α, for a given multi-index $\alpha = (\alpha_1, \dots, \alpha_k)$, of length $|\alpha| = k$. Whereas the integral

$$\int_{\mathbb{R}^d} \prod_{i=1}^k u_{\alpha_i} \phi(u)\, du, \quad 1 \le k \le s$$

is a product of integrals of the form

$$\int_{\mathbb{R}} u_j^{p_j} \phi_j(u_j)\, du_j, \quad \text{with} \quad j = 1, \dots, d \quad \text{et} \quad 1 \le p_j \le k \le s,$$

The constant C is given by

$$C = c_s \Big(\|f^{(s+1)}\|_\infty \Big) \int_{\mathbb{R}^d} \|u\|^{s+1} |\phi(u)|\, du$$

where c_s is a universal constant depending on s and $\|f^{(s+1)}\|_\infty$ is the sup norm of derivatives of order $s+1$ of f.

2. Let $\varphi : \mathbb{R}^d \to \mathbb{R}$ be a positive integrable and bounded function. Suppose that $\int_{\mathbb{R}^d} \varphi(x)dx = 1$. Let $\varphi_{h,x}(y) = \frac{1}{h^d}\varphi\Big(\frac{y-x}{h}\Big)$, then for every continuous and bounded function f we have

$$\lim_{h\to 0} \int_{\mathbb{R}^d} f(y)\varphi_{h,x}(y)\, dy = f(x).$$

Proof. We have that

$$\begin{aligned} \int_{\mathbb{R}^d} f(y)\phi_{h,x}(y)\, dy - f(x) &= \int_{\mathbb{R}^d} \phi_{h,x}(y)(f(y) - f(x))\, dy \\ &= \int_{\mathbb{R}^d} \phi(u)(f(x+uh) - f(x))\, du \end{aligned}$$

Using a Taylor serie expansion of order s we obtain

$$\begin{aligned} \int_{\mathbb{R}^d} f(y)\phi_{h,x}(y)\, dy - f(x) &= \sum_{k=1}^s \frac{h^k}{k!} \sum_{|\alpha|=k} \partial^\alpha f(x) \int_{\mathbb{R}^d} \prod_{i=1}^k u_{\alpha_i}\phi(u)\, du \\ &+ \frac{h^{s+1}}{s!} \sum_{|\alpha|=s+1} \int_{\mathbb{R}^d} \int_0^1 (1-\lambda)^s \partial^\alpha f(x + \lambda uh) \prod_{i=1}^{s+1} u_{\alpha_i}\phi(u)\, d\lambda\, du. \end{aligned}$$

Since $(\phi_j)_{j=1,\dots,d}$ are super kernels of order s, we conclude that for $1 \le p_j \le s-1$ we have

$$\int_{\mathbb{R}} u_j^{p_j} \phi_j(u_j)\, du_j = 0.$$

Consequently,

$$\int_{\mathbb{R}^d} f(y)\phi_{h,x}(y)\,dy - f(x) = \frac{h^s}{s!}\sum_{|\alpha|=s}\partial^\alpha f(x)\int_{\mathbb{R}^d}\prod_{i=1}^{s} u_{\alpha_i}\phi(u)\,du$$
$$+\frac{h^{s+1}}{s!}\sum_{|\alpha|=s+1}\int_{\mathbb{R}^d}\int_0^1 (1-\lambda)^s\partial^\alpha f(x+\lambda uh)\prod_{i=1}^{s+1} u_{\alpha_i}\phi(u)\,d\lambda\,du.$$

In the following we evaluate the remainder term.

$$\Big|\int_{\mathbb{R}^d}\int_0^1 (1-\lambda)^s\partial^\alpha f(x+\lambda uh)\prod_{i=1}^{s+1} u_{\alpha_i}\phi(u)\,d\lambda\,du\Big| \le \|f^{(s+1)}\|_\infty \int_{\mathbb{R}^d}\|u\|^{s+1}|\phi(u)|\,du.$$

According to property a) of Definition 3.1, the right side of the inequality is finite and therefore the result follows. The proof of the second assertion follows from the Lebesgue theorem. □

The main theorem of this section gives us an expansion of order 1 of the weak error in the approximation of the density of the hypoelliptic diffusion X.

Before this we study the error process in a form that will also be useful when studying the stable convergence problem.

The error process $U^n = (U_t^n)_{0\le t\le T}$, defined by

$$U_t^n = X_t - X_t^n,$$

satisfies the equation

$$dU_t^n = \sum_{j=0}^{q}(\dot{f}_{t,j}^n).(X_t - X_{\eta_n(t)}^n)\,dY_t^j,$$

where

$$\dot{f}_{t,j}^n = \int_0^1 \nabla f_j\Big(X_{\eta_n(t)}^n + \lambda(X_t - X_{\eta_n(t)}^n)\Big)\,d\lambda.$$

Therefore the equation satisfied by U^n can be written as:

$$U_t^n = \int_0^t \sum_{j=0}^{q}\dot{f}_{s,j}^n\,dY_s^j.U_s^n + G_t^n, \tag{6}$$

with

$$G_t^n = \int_0^t \sum_{j=0}^{q}\dot{f}_{s,j}^n.(X_s^n - X_{\eta_n(s)}^n)\,dY_s^j. \tag{7}$$

Note that

$$X_s^n - X_{\eta_n(s)}^n = \sum_{j=0}^{q}\bar{f}_{s,j}^n(Y_s^j - Y_{\eta_n(s)}^j), \tag{8}$$

with $\bar{f}_{s,j}^n = f_j(X_{\eta_n(s)}^n)$. In the following let $(Z_t^n)_{0\le t\le T}$ be the $\mathbb{R}^{d\times d}$ valued solution of

$$Z_t^n = I_d + \int_0^t \sum_{j=0}^{q}\dot{f}_{s,j}^n\,dY_s^j.Z_s^n\,.$$

From Theorem 56 p.271 in Protter (1990) we obtain that there exists $(Z_s^n)^{-1}$ for all $s\le T$ which satisfies

$$(Z_t^n)^{-1} = I_d - \int_0^t (Z_s^n)^{-1}\sum_{j=1}^{q}(\dot{f}_{s,j}^n)^2 ds - \int_0^t (Z_s^n)^{-1}\sum_{j=0}^{q}\dot{f}_{s,j}^n dY_s^j$$

and that

$$U_t^n = Z_t^n\Big\{\int_0^t (Z_s^n)^{-1} dG_s^n - \int_0^t (Z_s^n)^{-1}\sum_{j=1}^q (\dot f_{s,j}^n)^2 (X_s^n - X_{\eta_n(s)}^n)\, ds\Big\}.$$

We define $Z_t = D_x X_t$ and therefore we have that it satisfies

$$Z_t = I_d + \int_0^t \sum_{j=0}^q \dot f_{s,j}\, dY_s^j . Z_s\,.$$

with $\dot f_{t,j} = \nabla f_j(X_t)$.

Furthermore Z_t^{-1} exists and satisfies the following explicit linear stochastic differential equation

$$(Z_t)^{-1} = I_d - \int_0^t (Z_s)^{-1}\sum_{j=1}^q (\dot f_{s,j})^2 ds - \int_0^t (Z_s)^{-1}\sum_{j=0}^q \dot f_{s,j} dY_s^j$$

Then using the same technique as in the proof of existence and uniqueness for stochastic differential equations with Lipschitz coefficients (i.e. Gronwall inequality), we obtain that

$$\forall p\geq 1 \qquad \lim_{n\to\infty} \mathbb{E}\Big[\sup_{0\leq t\leq T} \|Z_t^n - Z_t\|^p\Big] = 0,$$

and

$$\forall p\geq 1 \qquad \lim_{n\to\infty} \mathbb{E}\Big[\sup_{0\leq t\leq T} \Big\|(Z_t^n)^{-1} - (Z_t)^{-1}\Big\|^p\Big] = 0,$$

Now we are ready to give the main theorem in this section.

Theorem 3.1. *Under the above notations,*

1. *Let $h = n^{-\alpha}$, $\alpha \geq 1/s$. Then there exists a constant $C_{\phi,x}^s > 0$ depending on ϕ, $p(x)$ and s such that*
$$\mathbb{E}\Big[\phi_{h,x}\big(X_T^n + Z_{n,\theta}\big)\Big] - p(x) = \frac{C_{\phi,x}^s}{n} + o\left(\frac{1}{n}\right). \tag{9}$$

2. *let $\varphi \in \mathcal{C}_b^\infty(\mathbb{R}^d;\mathbb{R})$ be a positive bounded and integrable function with bounded derivatives. Suppose that $\int_{\mathbb{R}^d}\varphi(x)dx = 1$. Let*
$$\varphi_{h,x}(y) = \frac{1}{h^d}\varphi\Big(\frac{y-x}{h}\Big), \quad h = n^{-\alpha} \quad \textit{with} \quad \alpha > 0,$$
then we have
$$\lim_{n\to 0} \mathbb{E}\,\varphi_{h,x}(X_T^n + Z_{n,\theta}) = p(x).$$

Proof. First we give the proof of the first assertion.

• Proof of the first assertion

We write the weak approximation error as follows

$$\begin{aligned}\mathbb{E}\Big[\phi_{h,x}\big(X_T^n + Z_{n,\theta}\big)\Big] - p(x) =& \mathbb{E}\Big[\phi_{h,x}\big(X_T^n + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big] \\ &+ \mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] \\ &+ \mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] - p(x).\end{aligned}$$

• Step 1:

We study the last term given by: $\mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] - p(x)$. In fact, using the regularity of the density of the diffusion X (under the Hörmander conditions), we obtain using the first assertion of the previous lemma that

$$\mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] - p(x) = \frac{h^s}{s!}\sum_{|\beta|=s}\partial^\beta p(x)\int_{\mathbb{R}^d}\prod_{i=1}^{s}u_{\beta_i}\phi(u)\,du + o(h^s),$$

where $\partial^\beta p$ is the partial derivative of p corresponding to the multi-index β. Note that for $h = n^{-\alpha}$, $\alpha \geq 1/s$ we have $o(h^s) = o(1/n)$.

• **Step 2:**

The second term is given by: $\mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big]$, we have

$$\begin{aligned}\mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] &= \frac{1}{2n^{1+2\theta}}\sum_{k=1}^{d}\mathbb{E}\big(\partial^2_{kk}\phi_{h,x}(X_T)\big)\\ &+ \frac{1}{3!}\,\mathbb{E}\int_0^1(1-\lambda)^3\big(Z_{n,\theta}.\nabla\big)^4\phi_{h,x}\big(X_T+\lambda Z_{n,\theta}\big)\,d\lambda.\end{aligned}$$

Using the integration by parts formula we obtain

$$\begin{aligned}\mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T\big)\Big] &= \frac{1}{2n^{1+2\theta}}\sum_{k=1}^{d}\partial^2_{kk}p(x) + \frac{1}{2n^{1+2\theta}}\sum_{k=1}^{d}\int_{\mathbb{R}^d}\phi_{h,x}(y)\partial^2_{kk}(p(y)-p(x))\,dy\\ &+ \frac{1}{3!}\,\mathbb{E}\int_0^1(1-\lambda)^3\big(Z_{n,\theta}.\nabla\big)^4\phi_{h,x}\big(X_T+\lambda Z_{n,\theta}\big)\,d\lambda.\end{aligned}$$

Thanks to the first assertion of the previous lemma we obtain

$$\int_{\mathbb{R}^d}\phi_{h,x}(y)\partial^2_{kk}(p(y)-p(x))\,dy = o(h^s) = o(1/n).$$

In addition, since $Z_{n,\theta}$ and X are independent we obtain, after applying the integration by parts formula four times, that

$$\begin{aligned}\mathbb{E}\Big[\big(Z_{n,\theta}.\nabla\big)^4\phi_{h,x}\big(X_T+\lambda Z_{n,\theta}\big)\Big] &= \mathbb{E}\int_{\mathbb{R}^d}\big(Z_{n,\theta}.\nabla\big)^4\phi_{h,x}\big(y+\lambda Z_{n,\theta}\big)p(y)\,dy\\ &= \mathbb{E}\int_{\mathbb{R}^d}\phi_{h,x}\big(y+\lambda Z_{n,\theta}\big)\big(Z_{n,\theta}.\nabla\big)^4p(y)\,dy\end{aligned}$$

Since $\phi_{h,x}$ is bounded we obtain that

$$\Big|\mathbb{E}\int_{\mathbb{R}^d}\phi_{h,x}\big(y+\lambda Z_{n,\theta}\big)\big(Z_{n,\theta}.\nabla\big)^4p(y)\,dy\Big| \leq \frac{c}{n^{4(1+\theta)d}}.$$

The last inequality is immediate using the definition of $Z_{n,\theta}$ and that $\nabla^4 p$ is integrable, since p decreases exponentially fast (see Kusuoka and Stroock (1985)) . The result follows.

• **Step 3:**

Now we deal with the first term given by

$$A_n = \mathbb{E}\Big[\phi_{h,x}\big(X_T^n + Z_{n,\theta}\big)\Big] - \mathbb{E}\Big[\phi_{h,x}\big(X_T + Z_{n,\theta}\big)\Big].$$

In fact, we have

$$A_n = \int_0^1\mathbb{E}\Big(\nabla\phi_{h,x}\big(\zeta_\lambda^n + Z_{n,\theta}\big).U_T^n\Big)d\lambda, \tag{10}$$

where $\zeta_\lambda^n = X_T + \lambda(X_T^n - X_T)$. In what follows we use the ideas contained in Clement et al. (2004). Recalling equations (6), (7) and (8) we have that

$$A_n = \sum_{j,k=0}^{q}E\Big(\int_0^1\nabla\phi_{h,x}\big(\zeta_\lambda^n + Z_{n,\theta}\big)d\lambda Z_T^n\int_0^T(Z_s^n)^{-1}F_{jk}^n(s)(Y_s^j - Y_{\eta(s)}^j)dY_s^k\Big)$$

where $F^n_{jk}(s) = \dot{f}^n_{s,j}\bar{f}^n_{s,k}$. If we define $D^0 = I$ (the identity operator) then using the duality formula (1), one obtains

$$A_n = \sum_{j,k=0}^{q} E\Big(\int_0^1\int_0^T\int_{\eta(s)}^{s} D^j_u\left\{D^k_s\{\nabla\phi_{h,x}(\zeta^n_\lambda + Z_{n,\theta})Z^n_T\}(Z^n_s)^{-1}F^n_{jk}(s)\right\}dudsd\lambda\Big).$$

Next, if we apply the stochastic derivative operators one obtains that the above is a sum of terms of the type

$$E\Big(\int_0^1\int_0^T\int_{\eta(s)}^{s}\partial^r\phi_{h,x}(\zeta^n_\lambda + Z_{n,\theta})G^{n,r,j,k}_{u,s}dudsd\lambda\Big), \tag{11}$$

where $j,\ k = 0,...,q$ and r is a multi-index of order 1 up to order 3. The random variables $G^{n,r,j,k}_{u,s}$ are given by

$$\begin{array}{l}
\left(D^j_u\left\{D^k_s\{Z^n_T\}\right\}(Z^n_s)^{-1}F^n_{jk}(s) + D^k_s\{Z^n_T\}D^j_u\left\{(Z^n_s)^{-1}F^n_{jk}(s)\right\}\right)^a \text{ if } r=(a)\\
(D^k_s\{\zeta^n_\lambda\})^a\left(D^j_u\{Z^n_T\}(Z^n_s)^{-1}F^n_{jk}(s)\right)^b + (D^k_s\{\zeta^n_\lambda\})^a\left(Z^n_T D^j_u\{(Z^n_s)^{-1}F^n_{jk}(s)\}\right)^b +\\
(D^j_u\{\zeta^n_\lambda\})^a\left(D^k_s\{Z^n_T\}(Z^n_s)^{-1}F^n_{jk}(s)\right)^b \text{ if } r=(a,b)\\
(D^j_u\{\zeta^n_\lambda\})^a(D^k_s\{\zeta^n_\lambda\})^b(Z^n_T(Z^n_s)^{-1}F^n_{jk}(s))^c \text{ if } r=(a,b,c).
\end{array}$$

Here $a,\ b,\ c \in \{1,...,d\}$ denote the component of the corresponding vector. Next for each term one applied the integration by parts formula (2) to obtain that each term of the type (11) can be written as

$$B_n(r,j,k) := E\Big(\int_0^1\int_0^T\int_{\eta(s)}^{s}\psi_{h,x}(\zeta^n_\lambda)\tilde{H}_{r+}(\zeta^n_\lambda + Z_{n,\theta}, G^{n,r,j,k}_{u,s})dudsd\lambda\Big),$$

where $r+ = (r,1,...,d)$ and $\psi_{h,x}(y) := \int_{(-\infty,y_i)^d}\phi_{h,x}(t)\,dt$

The proof of the first assertion follows using the following two lemmas which are proved in the appendix.

Lemma 3.2. *Let $g, g_n : [0,T]\times[0,T] \to \mathbb{R},\ \ n\in\mathbb{N}$. Suppose that*

i) g is continuous on the compact $[0,T]\times[0,T]$.

ii) $\sup_{0\le s,u\le T}|g_n(s,u) - g(s,u)| \underset{n\to\infty}{\longrightarrow} 0.$

Then

$$\int_0^T\int_{\eta_n(u)}^{u} g_n(s,u)\,ds\,du = \frac{1}{2n}\int_0^T g(u,u)\,du + o(1/n).$$

Lemma 3.3. *Under the previous notations we obtain*

$$B_n(r,j,k) = \frac{1}{2n}\int_0^T \mathbb{E}\Big(\mathbf{1}_{\{X_T>x\}}\,\mathbf{H}_{r^+}\left(X_T, G^{r,j,k}_u\right)\Big)\,du + o\left(\frac{1}{n}\right), \tag{12}$$

with $G^{r,j,k}_u$ is the limit process given by (here $F_{jk}(s) = \dot{f}_{s,j}f_k(X_t)$)

$$\begin{array}{l}
\left(D^j_u\left\{D^k_s\{Z_T\}\right\}(Z_s)^{-1}F_{jk}(s) + D^k_s\{Z_T\}D^j_u\left\{(Z_s)^{-1}F_{jk}(s)\right\}\right)^a \textit{ if } r=(a)\\
(D^k_s\{X_T\})^a\left(D^j_u\{Z_T\}(Z_s)^{-1}F_{jk}(s)\right)^b + (D^k_s\{X_T\})^a\left(Z_T D^j_u\{(Z_s)^{-1}F_{jk}(s)\}\right)^b +\\
(D^j_u\{X_T\})^a\left(D^k_s\{Z_T\}(Z_s)^{-1}F_{jk}(s)\right)^b \textit{ if } r=(a,b)\\
(D^j_u\{X_T\})^a(D^k_s\{X_T\})^b(Z_T(Z_s)^{-1}F_{jk}(s))^c \textit{ if } r=(a,b,c).
\end{array}$$

The proof of the second assertion follows as the first assertion with the exception that the rate is not $1/n$ but $1/n^{2\alpha}$ if $\alpha < 1/2$. We mention here that in the proof of the third step above we only need the integrability of ϕ and that $\int_{\mathbb{R}^d} \phi(x)dx = 1$. Consequently, the results obtained in this step remain valid in the context of the second assertion of the theorem. $\square$

4 Approximations of non-degenerated diffusions through the Monte Carlo method

Let X be a hypoelliptic diffusion solution of the stochastic differential equation (3). The goal of this section is to study an approximation of the density $p(x)$ of $X(T)$ using a Monte Carlo method together with a kernel density estimate. That is, in order to evaluate $p(x)$:

• One discretizes the diffusion X through an Euler scheme X^n of step T/n regularized as $X^n + Z_{n,\theta}$ where $Z_{n,\theta}$ is an independent Gaussian random variable of mean zero and standard deviation $n^{-1/2-\theta}$.
• one approximates the distribution $y \mapsto \delta_x(y)$ by the super-kernel $\phi_{h,x}(y)$ of order s , where h denotes the window size.
• then finally one estimates $\mathbb{E}\,\phi_{h,x}(X_T^n + Z_{n,\theta})$ using the Monte Carlo method. This procedure gives the classical kernel estimator given by

$$\hat{S}^{n,N} := \frac{1}{N}\sum_{i=1}^{N} \phi_{h,x}(X_{T,i}^n + Z_{n,\theta}^i)$$

where $(X_{T,i}^n)_{1\le i\le N}$ and $(Z_{n,\theta}^i)_{1\le i\le N}$ are i.i.d. copies of X_T^n and $Z_{n,\theta}$. In what follows, we prove a central limit theorem analogue to a similar result proved by Duffie and Glynn (1995) which gives a precise choice for the sample size N for the Monte Carlo method. This choice depends on the step size parameter n from the Euler scheme and is valid for the regular case. Here we extend this result to the degenerate case, the problem is somewhat more complex as we have to decide the optimal values of N and h in function of n.

In what follows we let $N = n^\gamma$, $h = n^{-\alpha}$ where $\gamma > 0$ and $\alpha \ge 1/s$

Theorem 4.1. *With the previous definitions and if we let $\gamma = 2 + \alpha d$ then*

$$n(S^{n,N} - p(x)) \Rightarrow \sigma G + C_{\phi,x}^s$$

with $\sigma^2 = \phi_2\, p(x)$, G is a standard Gaussian random variable and $C_{\phi,x}^s$ is the constant in the error expansion given in Theorem 3.1 and $\phi_2 = \int_{\mathbb{R}^d} |\phi(u)|^2\, du$.

Proof. We have that

$$n(S^{n,N} - p(x)) = \frac{1}{n^{\gamma-1}}\sum_{i=1}^{n^\gamma}\Big\{\phi_{h,x}(X_{T,i}^n + Z_{n,\theta}^i) - \mathbb{E}\big[\phi_{h,x}(X_T^n + Z_{n,\theta})\big]\Big\}$$
$$+\, n\Big[\mathbb{E}\big[\phi_{h,x}(X_T^n + Z_{n,\theta})\big] - p(x)\Big].$$

From Theorem 3.1, we have that

$$n\Big[\mathbb{E}\big[\phi_{h,x}(X_T^n + Z_{n,\theta})\big] - p(x)\Big] \underset{n\to\infty}{\longrightarrow} C_{\phi,x}^s$$

Therefore it remains to prove a central limit theorem for $\frac{1}{n^{\gamma-1}}\sum_{i=1}^{n^\gamma}\zeta_{T,i}^{n,h}$ where

$$\zeta_{T,i}^{n,h} := \Big\{\phi_{h,x}(X_{T,i}^n + Z_{n,\theta}^i) - \mathbb{E}\big[\phi_{h,x}(X_T^n + Z_{n,\theta})\big]\Big\}.$$

We start considering the characteristic function of the previous sum

$$\mathbb{E}\left[\exp\Big(\frac{iu}{n^{\gamma-1}}\sum_{k=1}^{n^\gamma}\zeta_{T,k}^{n,h}\Big)\right] = \left[\mathbb{E}\exp\Big(\frac{iu\zeta_T^{n,h}}{n^{\gamma-1}}\Big)\right]^{n^\gamma}$$
$$= \left[1+\frac{1}{n^\gamma}\Big(\frac{-u^2}{2n^{\gamma-2}}\mathbb{E}\,|\zeta_T^{n,h}|^2 + \mathbb{E}\,C_{n,h}(\omega)\Big)\right]^{n^\gamma}.$$

Here

$$|\mathbb{E}\,C_{n,h}(\omega)| \le \frac{u^3}{6n^{2\gamma-3}}\mathbb{E}\,|\zeta_T^{n,h}|^3.$$

To study the above terms we define the following kernels

$$\phi_{2,h,x}(y) = \frac{1}{h^d\phi_2}\phi^2\Big(\frac{y-x}{h}\Big), \qquad \phi_2 = \int_{\mathbb{R}^d}\phi^2(u)\,du$$

and

$$\phi_{3,h,x}(y) = \frac{1}{h^d\phi_3}\Big|\phi^3\Big(\frac{y-x}{h}\Big)\Big|, \qquad \phi_3 = \int_{\mathbb{R}^d}|\phi^3(u)|\,du.$$

These two positive functions are integrable and integrate to one. Therefore from the second assertion of Theorem 3.1 we have

$$\mathbb{E}\Big[\phi_{i,h,x}(X_T^n + Z_{n,\theta})\Big] = p(x) + \varepsilon_i(x), \quad i = 2,3.$$

with $\lim_n \varepsilon_i(x) = 0$ for $i = 2,3$.

Let's start studying the term given by $\mathbb{E}\,|\zeta_T^{n,h}|^2$. We have that

$$\mathbb{E}\,|\zeta_T^{n,h}|^2 = \mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta})^2\right] - \Big\{\mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta})\right]\Big\}^2$$
$$= \frac{\phi_2}{h^d}\mathbb{E}\left[\phi_{2,h,x}(X_T^n + Z_{n,\theta})\right] - \Big\{\mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta})\right]\Big\}^2.$$

Therefore,

$$\mathbb{E}\,|\zeta_T^{n,h}|^2 = \frac{\phi_2}{h^d}\varepsilon_2(x) + \frac{\phi_2}{h^d}p(x) + \Big\{\frac{C^s_{\phi,x}}{n} + o\Big(\frac{1}{n}\Big) + p(x)\Big\}^2$$

where $C^s_{\phi,x}$ is the constant in the error expansion given in Theorem 3.1. Therefore, for $h = n^{-\alpha}$, $\gamma = 2+\alpha d$ et $\alpha \ge 1/s$ we have

$$\frac{1}{n^{\gamma-2}}\mathbb{E}\,|\zeta_T^{n,h}|^2 \underset{n\to\infty}{\longrightarrow} \phi_2\,p(x).$$

On the other hand, we have that

$$\mathbb{E}\,|\zeta_T^{n,h}|^3 = \mathbb{E}\,\Big|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \mathbb{E}\big[\phi_{h,x}(X_T^n + Z_{n,\theta})\big]\Big|^3$$
$$\le \mathbb{E}\,\big|\phi_{h,x}(X_T^n + Z_{n,\theta})\big|^3 + 3\mathbb{E}\,\big|\phi_{h,x}(X_T^n + Z_{n,\theta})\big|^2\big|\mathbb{E}\,\phi_{h,x}(X_T^n + Z_{n,\theta})\big|$$
$$+ 4\big|\mathbb{E}\,\phi_{h,x}(X_T^n + Z_{n,\theta})\big|^3.$$

Therefore, as before, we obtain that

$$\mathbb{E}\,|\zeta_T^{n,h}|^3 \le h^{-2d}\phi_3\mathbb{E}\,\phi_{3,h,x}(X_T^n + Z_{n,\theta}) + 3h^{-d}\phi_2\mathbb{E}\,\phi_{2,h,x}(X_T^n + Z_{n,\theta})\big|\mathbb{E}\,\phi_{h,x}(X_T^n + Z_{n,\theta})\big|$$
$$+ 4\big|\mathbb{E}\,\phi_{h,x}(X_T^n + Z_{n,\theta})\big|^3$$

Finally,

$$\mathbb{E}\,|\zeta_T^{n,h}|^3 \leq h^{-2d}\phi_3(p(x)+\varepsilon_3(x)) + 3h^{-d}\phi_2\Big|p(x)+\varepsilon_2(x)\Big|\Big|\frac{C^s_{\phi,x}}{n}+o\Big(\frac{1}{n}\Big)+p(x)\Big|$$

$$+4\Big|\frac{C^s_{\phi,x}}{n}+o\Big(\frac{1}{n}\Big)+p(x)\Big|^3.$$

for $h=n^{-\alpha}$, $\gamma=2+\alpha d$ and $\alpha\geq 1/s$. This leads to

$$\frac{1}{n^{2\gamma-3}}\mathbb{E}\,|\zeta_T^{n,h}|^3 \underset{n\to\infty}{\longrightarrow} 0.$$

which finishes the proof. □

The interpretation of the above result leads to the previously announced result. That is, in order to approximate the density $p(x)$ through a Monte Carlo method with a tolerance error of order $1/n$, the optimal asymptotic choice of parameters are $h=n^{-\alpha}$ and $N=n^{2+\alpha d}$ with $\alpha\geq 1/s$ where s denotes the order of the super kernel used for the estimation. This leads to the following algorithmic complexity (that is, number of calculations) of

$$C_{MC} = C\times nN = C\times n^{3+\alpha d},$$

for a given $C>0$ (here the unit of calculation is one simulation of a random variable). Therefore the optimal complexity of this algorithm is given by

$$C^\star_{MC} = C\times n^{3+\frac{d}{s}}.$$

Therefore we conclude that if the order s of the kernel is bigger then the complexity is smaller.

Nevertheless, one should keep in mind that the constant $C^s_{\phi,p(x)}$ depends on s and the implementation of this algorithm for high order kernels carries some problems, such as non-positive estimates and big constants in the error expansions. Therefore the practical choice of super kernel remains an open problem from the practical point of view.

5 Asymptotic behaviour of the Malliavin derivative of the normalised error

5.1 Malliavin derivative of the error process

In the following we denote $\check{W}^n$ the $d\times d$-dimensional process defined by

$$\check{W}^{n,ij}_t = \sqrt{\frac{2n}{T}}\int_0^t (W^i_s - W^i_{\eta_n(s)})\,dW^j_s.$$

According to the theorem 3.2 of Jacod and Protter (1998), the process $\check{W}^n$ converge stably in law to a bi-dimensional Brownian motion $\check{W}$ independent from W and the couple $(\check{W}^n,\sqrt{n}U^n)$ converge stably in law to the couple $(\check{W},U)$ where the $\mathbb{R}^{d\times d}$-valued process U is solution to

$$U_t = \sum_{j=0}^q \int_0^t \dot{f}_{s,j}.U_s\,dY^j_s + \sqrt{\frac{T}{2}}\sum_{i,j=1}^q\int_0^t \dot{f}_{s,j}.f_i(X_s)\,d\check{W}^{ij}_s. \tag{13}$$

In order to obtain the equation satisfied by the Malliavin derivative of the error process with respect to W^i, $i=1,...,q$, we derive the equation (6):

$$D^i_sU^n_t = \sum_{j=0}^q\int_0^t D^i_s(\dot{f}^n_{j,v}.U^n_v)\,dY^j_v + \dot{f}^n_{s,i}U^n_s\mathbf{1}_{\{s\leq t\}} + D^i_sG^n_t. \tag{14}$$

Note that the above derivative exists due to the regularity properties of the coefficients of the equation for X. Furthermore, using (7) and (8), we have that

$$D_s^i G_t^n = \dot{f}_{s,i}^n.(X_s^n - X_{\eta_n(s)}^n)\mathbf{1}_{\{s\le t\}} + + \sum_{j=0}^{q}\int_0^t D_s^i\Big[\dot{f}_{u,j}^n.(X_u^n - X_{\eta_n(u)}^n)\Big]\,dY_u^j$$

and

$$D_s^i\Big[\dot{f}_{u,j}^n.(X_u^n - X_{\eta_n(u)}^n)\Big] = \sum_{k=0}^{q} D_s^i(\dot{f}_{u,j}^n.\bar{f}_{u,k}^n)(Y_u^k - Y_{\eta_n(u)}^k) + \dot{f}_{u,j}^n.\bar{f}_{u,i}^n\mathbf{1}_{\{\eta_n(u)\le s\le u\}}$$

As $D_s Z = 0$ for Z, which is $\mathcal{F}_u$-measurable ($u < s$), the relation (14) becomes for $s \le t$,

$$D_s^i U_t^n = \dot{f}_{s,i}^n(U_s^n + X_s^n - X_{\eta_n(s)}^n) + \sum_{j=0}^{q}\int_s^t \dot{f}_{v,j}^n D_s^i U_v^n\,dW_v^j + \tilde{G}_{s,t}^{n,i}, \tag{15}$$

with

$$\tilde{G}_{s,t}^{n,i} = \sum_{j=0}^{q}\int_s^t D_s^i\dot{f}_{v,j}^n\,U_v^n\,dY_v^j + \sum_{j,k=0}^{q}\int_s^t D_s^i(\dot{f}_{u,j}^n\,\bar{f}_{u,k}^n)(Y_u^k - Y_{\eta_n(u)}^k)\,dY_u^j$$
$$+\sum_{j=0}^{q}\int_s^t \dot{f}_{u,j}^n\,\bar{f}_{u,i}^n\mathbf{1}_{\{\eta_n(u)\le s\le u\}}\,dY_u^j. \tag{16}$$

From Theorem 56 p.271 in Protter (1990), it follows that (15) becomes for $t \ge s$,

$$D_s^i U_t^n = Z_t^n(Z_s^n)^{-1}\dot{f}_{s,i}^n(U_s^n + X_s^n - X_{\eta_n(s)}^n)$$
$$+ Z_t^n\Big\{\int_s^t (Z_u^n)^{-1}\,d\tilde{G}_{s,u}^{n,i} - \sum_{j=0}^{q}\int_s^t (Z_u^n)^{-1}\dot{f}_{u,j}^n d\langle \tilde{G}_{s,.}^{n,i}, Y^j\rangle_u\Big\}. \tag{17}$$

5.2 A law convergence theorem for the normalised Malliavin derivative

The Malliavin derivative of U_T^n is a random vector taking values in the Hilbert space $H = L^2([0,T])$. The aim of this section is to establish the convergence in law for the sequence $\sqrt{n}DU_T^n$. Note that the process U, limit of $\sqrt{n}U^n$, is an adapted process with respect to the filtration of W and $\check{W}$. Using (13), we can compute the derivatives DU_t and $\check{D}U_t$ with respect to both Wiener processes W and $\check{W}$ to obtain that DU_t satisfies for $0 \le s \le t \le T$,

$$D_s^i U_t = \dot{f}_{s,i}U_s + \sum_{j=0}^{q}\int_s^t \dot{f}_{v,j}D_s^i U_v\,dY_v^j + \sum_{j=0}^{q}\int_s^t D_s^i\dot{f}_{v,i}U_v\,dY_v^j$$
$$+\sqrt{\frac{T}{2}}\sum_{j,k=1}^{q}\int_s^t D_s^i(\dot{f}_{v,j}f_{v,k})d\check{W}_v^{kj}, \tag{18}$$

or using again Theorem 56 p.271 in Protter (1990), we obtain for $0 \le s \le t \le T$ that,

$$D_s U_t = Z_t(Z_s)^{-1}\dot{\sigma}_{s,i}U_s + Z_t\Big\{\int_s^t (Z_u)^{-1}dG_{s,u}^i - \sum_{j=0}^{q}\int_s^t (Z_u)^{-1}\dot{f}_{u,j}\,d\langle G_{s,.}^i, Y^j\rangle_u\Big\}, \tag{19}$$

with

$$G_{s,t}^i = \sum_{j=0}^{q}\int_s^t D_s^i\dot{f}_{v,j}U_v\,dY_v^j + \sqrt{\frac{T}{2}}\sum_{j,k=1}^{q}\int_s^t D_s^i(\dot{f}_{v,j}f_{v,i})\,d\check{W}_v^{kj}. \tag{20}$$

Theorem 5.1. *Let $(H_t^i)_{0\le t\le T}$ be a continuous sequence of $\mathbb{R}$-valued process (possibly non adapted). The random vector $(\sqrt{n}U_T^n, \sqrt{n}\int_0^T H_s^i D_s^i U_T^n\, ds)$ converges stably in law to $(U_T, \int_0^T H_s^i D_s^i U_T\, ds)$ where $D^i U_T$ is the Malliavin derivative of U with respect to W^i and solution of (19).*

In order to prove this theorem, we use the two technical lemmas below. The proofs of these lemmas are given in the appendix. (See Jacod and Protter (1998) for related results).

Lemma 5.1. *Let $(H_t^n = (H_t^{1,n}, ..., H_t^{d,n}))_{0\le t\le T}$ be a sequence of continuous tight sequence of process (possibly non adapted) taking values in $\mathbb{R}^d$. The sequence of random vectors $(\sqrt{n}\int_0^T H_s^{i,n}(Y_s^j - Y_{\eta_n(s)}^j)\, ds; i\in\{1,...,d\}, j\in\{0,...,q\})_{n\in\mathbb{N}}$ converge in probability to 0.*

Lemma 5.2. *Let $(H_t)_{0\le t\le T}$ be a continuous $\mathbb{R}$-valued process (possibly non adapted) and let $(K_u^n)_{0\le u\le T}$ be a sequence of adapted and continuous processes taking values in $\mathbb{R}^d$ and such that $\sup_n \mathbb{E}\int_0^T \|K_u^n\|^2\, du < \infty$. Then the sequence $(\sqrt{n}\int_0^T H_s(\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}} K_u^n\, dW_u^j)ds)_{n\in\mathbb{N}}$ converge in probability to 0.*

In the following we denote

$$\bar{U}_t^n = \sqrt{n}U_t^n.$$

Lemma 5.3. *Let H^i, K^i, L^i be three real processes with continuous processes on $[0,T]$ and let $(\xi_{s,u}^{ij})_{0\le s\le u\le T}$, $(\zeta_{s,u}^{ijk})_{0\le s\le u\le T}$ be two processes, taking values in $\mathbb{R}^{d\times d}$, with continuous trajectories and such that*

$$\mathbb{E}\int_0^T du\int_0^u ds\Big(\max_j \|\xi_{s,u}^{ij}\|^p + \max_{j,k}\|\zeta_{s,u}^{ijk}\|^p\Big) < \infty \qquad \textit{for} \quad p>2,\ i=1,...,q.$$

Then

$$\Big(\bar{U}_T^n, \int_0^T H_s^i \bar{U}_s^n\, ds, \int_0^T K_s^i\Big(\sum_{j=1}^q\int_s^T \xi_{s,u}^{ij}\bar{U}_u^n\, dW_u^j\Big)ds,$$
$$\sqrt{n}\int_0^T L_s^i\Big(\sum_{j,k=1}^q \int_s^T \zeta_{s,u}^{ijk}\, d\check{W}_u^{n,kj}\Big)ds; i=1,...,q\Big)$$

stably converge in law to

$$\Big(U_T, \int_0^T H_s^i U_s\, ds, \int_0^T K_s^i\Big(\sum_{j=1}^q\int_s^T U_u\xi_{s,u}^{ij}\, dW_u^j\Big)ds,$$
$$\int_0^T L_s^i\Big(\sum_{j,k=1}^q \int_s^T \zeta_{s,u}^{ijk}\, d\check{W}_u^{kj}\Big)ds, i=1,...,q\Big)$$

Proof of Theorem 5.1. Using the relation (17), we have

$$\begin{aligned} D_s^i U_T^n &= Z_T^n (Z_s^n)^{-1}\dot{f}_{s,i}^n(U_s^n + X_s^n - X_{\eta_n(s)}^n) \\ &\quad + Z_T^n\Big\{\int_s^T (Z_u^n)^{-1}\, d\tilde{G}_{s,u}^{n,i} - \sum_{j=0}^q\int_s^T (Z_u^n)^{-1}\dot{f}_{u,j}^n d\langle \tilde{G}_{s,.}^{n,i}, Y^j\rangle_u\Big\}. \end{aligned} \tag{21}$$

Consequently,

$$\int_0^T H_s^i D_s^i U_T^n\, ds = Z_T^n\int_0^T H_s^i (Z_s^n)^{-1}\dot{f}_{s,i}^n(U_s^n + X_s^n - X_{\eta_n(s)}^n)\, ds + Z_T^n I_T^{n,i}, \tag{22}$$

where

$$I_T^{n,i} = \int_0^T H_s^i\Big(\int_s^T (Z_u^n)^{-1} d\tilde{G}_{s,u}^{n,i} - \sum_{j=0}^q \int_s^T (Z_u^n)^{-1}\dot{f}_{u,j}^n\, d\langle\tilde{G}_{s,.}^{n,i}, Y^j\rangle_u\Big)\, ds.$$

Using (19)

$$\int_0^T H_s^i D_s^i U_T\, ds = Z_T \int_0^T H_s^i (Z_s)^{-1} \dot{f}_{s,i} U_s\, ds + Z_T I_T^i \tag{23}$$

with

$$I_T^i = \int_0^T H_s^i \Big(\int_s^T (Z_u)^{-1} dG_{s,u}^i - \sum_{j=0}^q \int_s^T (Z_u)^{-1} \dot{f}_{u,j}\, d\langle G_{s,.}^i, Y^j \rangle_u \Big)\, ds.$$

Note that

$$\int_0^T H_s^i (Z_s^n)^{-1} \dot{f}_{s,i}^n (U_s^n + X_s^n - X_{\eta_n(s)}^n)\, ds = \int_0^T H_s^i (Z_s)^{-1} \dot{f}_{s,i} U_s^n\, ds + \xi_T^{n,i} \tag{24}$$

with $\mathbb{P}\lim_{n\to\infty} \big(\sqrt{n}\xi_T^{n,i}\big) = 0$, where we use the notation $\mathbb{P}\lim$ for probability convergence.

In fact, the tightness of $\sqrt{n}U^n$ (see theorem 3.2 of Jacod and Protter (1998))and the convergence in probability of $\limsup_{0\le s\le T} |(Z_s^n)^{-1}\dot{f}_{s,j}^n - (Z_s)^{-1}\dot{f}_{s,j}|$ to 0 give that

$$\mathbb{P}\lim_{n\to\infty} \sqrt{n} \int_0^T H_s^i \big[(Z_s^n)^{-1} \dot{f}_{s,i}^n - (Z_s)^{-1} \dot{f}_{s,i} \big] U_s^n\, ds = 0.$$

In the other hand, we can write

$$\int_0^T H_s^i (Z_s^n)^{-1} \dot{f}_{s,i}^n (X_s^n - X_{\eta_n(s)}^n)\, ds = \sum_{j=0}^q \int_0^T H_s^i (Z_s^n)^{-1} \dot{f}_{s,i}^n \bar{f}_{s,j}^n (Y_s^j - Y_{\eta_n(s)}^j)\, ds.$$

We note that

$$\Big\| \sum_{j=0}^q \int_0^T H_s^i \big[(Z_s^n)^{-1} \dot{f}_{s,i}^n \bar{f}_{s,j}^n - (Z_s)^{-1} \dot{f}_{s,i} f_j(X_s) \big] (Y_s^j - Y_{\eta_n(s)}^j)\, ds \Big\| \le$$

$$\sum_{j=0}^q \sup_{0\le s\le T} \Big\| H_s^i \big[(Z_s^n)^{-1} \dot{f}_{s,i}^n \bar{f}_{s,j}^n - (Z_s)^{-1} \dot{f}_{s,i} f_j(X_s) \big] \Big\| \int_0^T |Y_s^j - Y_{\eta_n(s)}^j|\, ds$$

and using that the sequence $\sqrt{n}\int_0^T |Y_s^j - Y_{\eta_n(s)}^j|\, ds$ is tight (since bounded in L^1), it follows from lemma 5.1 that

$$\mathbb{P}\lim_{n\to\infty} \sqrt{n} \sum_{j=1}^q \int_0^T H_s (Z_s^n)^{-1} \dot{f}_{s,i}^n \bar{f}_{s,j}^n (Y_s^j - Y_{\eta_n(s)}^j)\, ds = 0.$$

Let's study now the sequence (I_T^n). First, note that using (16) we obtain that

$$I_T^{n,i} = \sum_{j=0}^q \int_0^T H_s^i \int_s^T (Z_u^n)^{-1} \Big\{ A_{s,u}^{n,i,j} U_u^n + \sum_{k=0}^q B_{s,u}^{n,i,j,k} (Y_u^k - Y_{\eta(u)}^k) + C_u^{n,i,j} \mathbf{1}_{\{\eta_n(u)\le s\le u\}} \Big\} dY_u^j ds \tag{25}$$

where

$$\begin{aligned}
A_{s,u}^{n,i,j} &= D_s^i \dot{f}_{u,j}^n - \mathbf{1}_{\{j=0\}} \sum_{l=1}^q \dot{f}_{u,l}^n D_s^i \dot{f}_{u,l}^n \\
B_{s,u}^{n,i,j,k} &= D_s^i (\dot{f}_{u,j}^n \bar{f}_{u,k}^n) - \mathbf{1}_{\{k=0\}} \sum_{l=1}^q \dot{f}_{u,l}^n D_s^i (\dot{f}_{u,l}^n \bar{f}_{u,k}^n) \\
C_u^{n,i,j} &= \dot{f}_{u,j}^n \bar{f}_{u,i}^n - \mathbf{1}_{\{j=0\}} \sum_{l=1}^q \dot{f}_{u,l}^n \dot{f}_{u,l}^n \bar{f}_{u,i}^n.
\end{aligned}$$

Now we study each of the three terms in (25). First, the third term in (25) satisfies that

$$\sum_{j=0}^{q}\sqrt{n}\int_0^T H_s^i\int_s^T (Z_u^n)^{-1}C_u^{n,i,j}\mathbf{1}_{\{\eta_n(u)\le s\le u\}}dY_u^j ds$$

tends to zero due to Lemma 5.2. Now, consider the second term.

$$\sum_{j,k=0;jk=0}^{q}\sqrt{n}\int_0^T H_s^i\int_s^T (Z_u^n)^{-1}B_{s,u}^{n,i,j,k}(Y_u^k-Y_{\eta(u)}^k)dY_u^j ds$$

tends to zero. Next, if we define $B_{s,u}^{i,j,k}=D_s^i(\dot{f}_{u,j}f_{u,k})$

$$\sum_{j,k=1}^{q}\sqrt{n}\int_0^T H_s^i\int_s^T (Z_u^n)^{-1}(B_{s,u}^{n,i,j,k}-B_{s,u}^{i,j,k})(Y_u^k-Y_{\eta(u)}^k)dY_u^j ds$$

tends to zero in $L^1(\Omega)$ as $\check{W}^{n,kj}$ is bounded uniformly in $L^p(\Omega)$ and $B_{s,u}^{n,i,j,k}-B_{s,u}^{i,j,k}$ converges to zero in $L^p(\Omega\times[0,T]^2)$, therefore this term also converges to zero. Then for the remaining

$$\sum_{j,k=1}^{q}\sqrt{\frac{T}{2}}\int_0^T H_s^i\int_s^T (Z_u^n)^{-1}B_{s,u}^{i,j,k}d\check{W}_u^{n,kj}ds,$$

we will apply Lemma 5.3 at the end together with the analysis for the first term of (25). For that first term, consider as previously

$$\sum_{j=0}^{q}\sqrt{n}\int_0^T H_s^i\int_s^T (Z_u^n)^{-1}(A_{s,u}^{n,i,j}-A_{s,u}^{i,j})U_u^n dY_u^j ds,$$

where $A_{s,u}^{i,j}=D_s^i\dot{f}_{u,j}-\mathbf{1}_{\{j=0\}}\sum_{l=1}^q \dot{f}_{u,l}D_s^i\dot{f}_{u,l}$. Again this term goes to zero in $L^1(\Omega)$ as the sequence $\sqrt{n}U^n$ is bounded uniformly in $L^p(\Omega)$ and $A_{s,u}^{n,i,j}-A_{s,u}^{i,j}$ converges to zero in $L^p(\Omega\times[0,T]^2)$. For the remaining term one applies together with the previous term, Lemma 5.3. □

6 An optimal control variate method for density estimation

The aim of this section, is to analyze the statistical Romberg method as a control variate introduced in Kebaier (2005) in the case of density estimation. In order to reduce variance in the density estimation of a non-degenerate d-dimensional diffusion $(X_t)_{0\le t\le T}$, we will use another estimation of the same density using less steps and simulation paths.

That is, we discretize the diffusion by two Euler schemes with time steps T/n and T/m $(m << n)$. Under the Hörmander condition, the statistical Romberg method approximates the density $p(x)$ of the diffusion $(X_t)_{0\le t\le T}$ by

$$\frac{1}{N_m}\sum_{i=1}^{N_m}\phi_{h,x}(\hat{X}_{T,i}^m+\hat{Z}_{m,\theta}^i)+\frac{1}{N_{n,m}}\sum_{i=1}^{N_{n,m}}\left\{\phi_{h,x}(X_{T,i}^n+Z_{n,\theta}^i)-\phi_{h,x}(X_{T,i}^m+Z_{m,\theta}^i)\right\},$$

where $\hat{X}_T^m$ is a second Euler scheme with step T/m and such that the Brownian paths used for X_T^n and X_T^m have to be independent of the Brownian paths used in order to simulate $\hat{X}_T^m$. Furthermore

$$Z_{n,\theta}=\frac{\tilde{W}_T}{n^{\frac{1}{2}+\theta}},\quad Z_{m,\theta}=\frac{\hat{W}_T}{m^{\frac{1}{2}+\theta}},\quad \theta\ge 0,$$

where $\hat{W}$ is a d-dimensional Brownian motion independent of W and $\tilde{W}$.

In order to run the statistical Romberg algorithm, we have to optimize the parameters in the method. In the same manner as in Kebaier (2005), we establish a central limit theorem which will lead to a precise description of how to choose the parameters N_m, $N_{n,m}$, m and h as functions of n. The essential difference with the problem studied in Kebaier (2005) is that the variance of the estimators explode. This issue will be resolved through an appropiate renormalization procedure and an appropiate decomposition of the derivatives of the kernel function.

In the following, we suppose that for a given $0 < \beta < 2/3$ we have

$$m = n^{\beta},\ N_n = n^{\gamma_1},\ \ N_{n,m} = n^{\gamma_2},\ \ h = n^{-\alpha},$$

where $\gamma_1, \gamma_2 > 0$, and $\alpha \geq 1/s$ (the parameter s denotes the order of the super-kernel ϕ). We set

$$V_n := \frac{1}{n^{\gamma_1}} \sum_{i=1}^{n^{\gamma_1}} \phi_{h,x}(\hat{X}^{n^{\beta}}_{T,i} + \hat{Z}^i_{n^{\beta},\theta}) + \frac{1}{n^{\gamma_2}} \sum_{i=1}^{n^{\gamma_2}} \left\{ \phi_{h,x}(X^n_{T,i} + Z^i_{n,\theta}) - \phi_{h,x}(X^{n^{\beta}}_{T,i} + Z^i_{n^{\beta},\theta}) \right\}.$$

Theorem 6.1. *Suppose that the first derivatives kernel function ϕ have the following decomposition*

$$\frac{\partial \phi}{\partial x_i}(x) = \phi_{1i}(x) - \phi_{2i}(x)$$

with

$$\phi_{ji} \geq 0 \quad \textit{and} \quad \int_{\mathbb{R}^d} |\phi_{ji}(x)|^2 dx < +\infty, \qquad \textit{for} \quad i = 1, ..., d,\ \ j = 1, 2.$$

Define

$$C_{ii'jj'} = \int_{\mathbb{R}^d} \phi_{ji}(x)\phi_{j'i'}(x)\, dx.$$

Let

$$\tilde{\sigma}^2 = \sum_{j,j'=1}^{2} \sum_{i,i'=1}^{d} C_{ii'jj'}(-1)^{j+j'} \left\{ \mathbb{E}\big[\delta_x(X_T)U^i_T U^{i'}_T\big] + T\delta_{ii'}p(x)\mathbf{1}_{\{\theta=0\}} \right\},$$

where $\delta_x(.)$ stands for the Dirac delta function and $\delta_{ii'}$ is the Kroeneker delta function. Assume that $h = n^{-\alpha}$, $\gamma_1 = 2 + \alpha d$, $\gamma_2 = (d+2)\alpha + 2 - \beta$ and $1/s \leq \alpha < \beta/(d+2)$ with $0 < \beta < 2/3$.

Then

$$n\big(V_n - p(x)\big) \Rightarrow \tilde{\sigma} G + C^s_{\phi,x}$$

where G is a standard Gaussian and $C^s_{\phi,x}$ is the discretization constant of Theorem 3.1.

Before proving this lemma we introduce an essential result about the rate of explosion of the variances of the estimators. In what follows we extend the previous notation to $\phi_{ji,h,x}(y) = \phi_{ji}(\frac{y-x}{h})$.

Lemma 6.1. *Under the notation and assumptions of the above theorem, we have*

1. $n^{\beta - \alpha(2+d)} \mathbb{E}\left[\phi_{h,x}(X^{n^{\beta}}_T + Z_{n^{\beta},\theta}) - \phi_{h,x}(X_T)\right]^2 \underset{n\to\infty}{\longrightarrow} \tilde{\sigma}^2.$

2. $n^{\beta - \alpha(2+d)} \mathbb{E}\left[\phi_{h,x}(X^{n}_T + Z_{n,\theta}) - \phi_{h,x}(X^{n^{\beta}}_T + Z_{n^{\beta},\theta})\right]^2 \underset{n\to\infty}{\longrightarrow} \tilde{\sigma}^2.$

We remark here that the assertion 1 above is satisfied also for $\beta \geq 2/3$.

Proof. Let's prove the first assertion of the lemma.

• **Step 1:**

The Taylor formula gives

$$\phi_{h,x}(X_T^{n^\beta}+Z_{n^\beta,\theta})-\phi_{h,x}(X_T)=\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)$$
$$+\frac{1}{2}\sum_{i,i'=1}^{d}\frac{\partial^2\phi_{h,x}}{\partial x_i\partial x_{i'}}(\xi_T^{n,ii'})\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\big(U_T^{n^\beta,i'}+Z^{i'}_{n^\beta,\theta}\big)$$

where $U_T^{n^\beta}=X_T^{n^\beta}-X_T$ and $\xi_T^n\in\prod_{i=1}^d[X_T^i,X_T^{n^\beta,i}+Z^i_{n^\beta,\theta}]$. Note that

$$\Big\|\frac{\partial^2\phi_{h,x}}{\partial x_i\partial x_{i'}}\Big\|_\infty\le h^{-(d+2)}\|\phi''\|_\infty,$$

where $\|\phi''\|_\infty=\max_{i,i'}\sup_{x\in\mathbb{R}^d}\Big|\frac{\partial^2\phi(x)}{\partial x_i\partial x_{i'}}\Big|$. Then there exists a constant $C_T>0$ such that

$$n^{\frac{\beta-\alpha(2+d)}{2}}\Big\|\frac{\partial^2\phi_{h,x}}{\partial x_i\partial x_{i'}}(\xi_T^{n,ii'})\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\big(U_T^{n^\beta,i'}+Z^{i'}_{n^\beta,\theta}\big)\Big\|_{L^2(\Omega)}\le C_T n^{\frac{-\alpha(2+d)-\beta}{2}}h^{-(d+2)}\|\phi''\|_\infty$$
$$=C_Tn^{\frac{\alpha(2+d)-\beta}{2}}\|\phi''\|_\infty\longrightarrow 0\quad\text{as}\quad n\to\infty.$$

Consequently, in order to obtain the first assertion of the lemma it suffices to prove that

$$n^{\frac{\beta-\alpha(2+d)}{2}}\Big\|\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\Big\|_{L^2(\Omega)}\longrightarrow\tilde\sigma\quad\text{as}\quad n\to\infty.\tag{26}$$

- **Step 2:** We have

$$\Big\|\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\Big\|^2_{L^2(\Omega)}=\sum_{i,i'=1}^{d}\mathbb{E}\Big\{\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\frac{\partial\phi_{h,x}}{\partial x_{i'}}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\big(U_T^{n^\beta,i'}+Z^{i'}_{n^\beta,\theta}\big)\Big\}$$
$$=\sum_{j,j'=1}^{2}\sum_{i,i'=1}^{d}\mathbb{E}\Big\{\frac{(-1)^{j+j'}}{h^{2+d}}\phi_{ji,h,x}(X_T)\phi_{j'i',h,x}(X_T)Y^{n^\beta}_{ii',\theta}\Big\},$$

where $Y^{n^\beta}_{ii',\theta}:=\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\big(U_T^{n^\beta,i'}+Z^{i'}_{n^\beta,\theta}\big)$ and

$$\frac{\partial\phi_{h,x}}{\partial x_i}(y)=\frac{1}{h}\big[\phi_{1i,h,x}(y)-\phi_{2i,h,x}(y)\big].$$

Then

$$\Big\|\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\Big\|^2_{L^2(\Omega)}=\sum_{j,j'=1}^{2}\sum_{i,i'=1}^{d}C_{ii'jj'}\frac{(-1)^{j+j'}}{h^{2+d}}\mathbb{E}\Big[\varphi^{ii'jj'}_{h,x}(X_T)Y^{n^\beta}_{ii',\theta}\Big],$$

where $\varphi^{ii'jj'}_{h,x}(y)=(C_{ii'jj'})^{-1}h^d\phi_{ji,h,x}(y)\phi_{j'i',h,x}(y)$. Let $(\xi_h^{ii'jj'})_{\{i,i'=1\ldots d,\ jj'=1,2\}}$ be i.i.d random vectors independent of all other random variables and also between themselves so that their density is given by $\varphi^{ii'jj'}_{h,x}(.)$. Without loss of generality, we assume that

$$\xi_h^{ii'jj'}\to 0\ \text{ a.s.}\quad\text{as } h\to 0.$$

Then

$$n^{\beta-\alpha(2+d)}\Big\|\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\Big\|^2_{L^2(\Omega)}$$
$$=n^\beta\sum_{j,j'=1}^{2}\sum_{i,i'=1}^{d}C_{ii'jj'}(-1)^{j+j'}\mathbb{E}\Big[\delta_x\big(X_T+\xi_h^{ii'jj'}\big)Y^{n^\beta}_{ii',\theta}\Big],$$

Applying the integration by parts formula of Malliavin-Thalmaier (see Theorem 4.23 in Malliavin and Thalmaier (2006)), we have

$$n^{\beta-\alpha(2+d)}\left\|\sum_{i=1}^{d}\frac{\partial\phi_{h,x}}{\partial x_i}(X_T)\big(U_T^{n^\beta,i}+Z^i_{n^\beta,\theta}\big)\right\|^2_{L^2(\Omega)}$$
$$=n^\beta\sum_{r=1}^{d}\sum_{j,j'=1}^{2}\sum_{i,i'=1}^{d}C_{ii'jj'}(-1)^{j+j'}\mathbb{E}\Big[\partial_r Q_d\big(X_T+\xi_h^{ii'jj'}-x\big)\,\mathbf{H}_{(r)}\big(X_T,Y^{n^\beta}_{ii',\theta}\big)\Big],$$

where Q is the fundamental solution of the Poisson equation in the following sense. If Δ denotes the Laplace operator and f is some function, then the solution of the equation $\Delta u=f$ is given by the convolution Q_d*f. The explicit expressions for Q_d are $Q_1(x)=x_+$, $Q_2(x)=a_2\ln|x|$ and $Q_d(x)=a_d|x|^{-(d-2)}$ for $d>2$ and suitable constants a_d, $d\geq 2$. . To deal with the last obtained quantity, we need the following technical lemma which is proven in the Appendix.

Lemma 6.2. *Let $(\xi_h^{ii'jj'})_{\{i,i'=1\ldots d,\ jj'=1,2\}}$ be i.i.d random vectors independent of all other random variables and also between themselves so that*

$$\xi_h^{ii'jj'}\to 0\quad a.s.\quad \text{as } h\to 0.$$

Then for $r=1,...,d$

1. $\partial_r Q_d\big(X_T+\xi_h^{ii'jj'}\big)\to\partial_r Q_d\big(X_T\big)$ *a.s.* *as* $h\to 0$.
2. *For any* $0<\delta<(d-1)^{-1}$, *we have* $\sup_{h>0}\mathbb{E}\Big|\partial_r Q_d\big(X_T+\xi_h^{ii'jj'}-x\big)\Big|^{1+\delta}<\infty$.

As the diffusion X and the associated Euler scheme satisfies Proposition 3.1 and using Proposition 7.1 then a

$$n^\beta\sup_n\Big\|\mathbf{H}_{(r)}\big(X_T,Y^{n^\beta}_{ii',\theta}\big)\Big\|_{k,p}<\infty. \tag{27}$$

Therefore, using classical convergence results and according to the Lemma 6.2, we only need to study the behaviour of

$$n^\beta\mathbb{E}\Big[\partial_r Q_d\big(X_T-x\big)\,\mathbf{H}_{(r)}\big(X_T,Y^{n^\beta}_{ii',\theta}\big)\Big]$$

in order to prove the relation (26).

We define the limit of $Y^{n^\beta}_{ii',\theta}$ as $Y_{ii',\theta}=(U_T^i+\hat{W}_T^i\mathbf{1}_{\{\theta=0\}})(U_T^{i'}+\hat{W}_T^{i'}\mathbf{1}_{\{\theta=0\}})$

• **Step 3:**

We have that

$$\mathbf{H}_{(r)}\big(X_T,Y^{n^\beta}_{ii',\theta}\big)=Y^{n^\beta}_{ii',\theta}\,\mathbf{H}_{(r)}(X_T,1)-\sum_{j,k=1}^{d}(\gamma^{-1}_{X_T})_{jr}\int_0^T D_s^kX_T^jD_s^kY^{n^\beta}_{ii',\theta}\,ds$$

Therefore as $s\mapsto D_sX_T$ is continuous for $s\in[0,T]$, we have due to Theorem 5.1 that

$$n^\beta\,\mathbf{H}_{(r)}\Big(X_T,Y^{n^\beta}_{ii',\theta}\Big)\Rightarrow^{stably}$$
$$Y_{ii',\theta}\,\mathbf{H}_{(r)}(X_T,1)-2\sum_{j,k=1}^{d}(\gamma^{-1}_{X_T})_{jr}\int_0^T D_s^k(U_T^iU_T^{i'})D_s^kX_T^j\,ds$$
$$=\mathbf{H}_{(r)}\Big(X_T,Y_{ii',\theta}\Big).$$

Due to (27) the sequence $n^\beta\,\mathbf{H}_{(r)}\Big(X_T, Y^{n^\beta}_{ii',\theta}\Big)$ is uniformly integrable and therefore

$$n^\beta \sum_{r=1}^d \mathbb{E}\Big\{\partial_r Q_d(X_T - x)\,\mathbf{H}_{(r)}\Big(X_T, Y^{n^\beta}_{ii',\theta}\Big)\Big\}$$
$$\underset{n\to\infty}{\longrightarrow} \sum_{r=1}^d \mathbb{E}\Bigg\{Q_d(X_T - x)\,\mathbf{H}_{(r)}\Big(X_T, Y_{ii',\theta}\Big)\Bigg\} = \mathbb{E}\Big\{\delta_x(X_T)Y_{ii',\theta}\Big\}.$$

The last equality follows from an application of the integration by parts formula. As $\hat{W}_T$ is independent from W, we have that

$$\mathbb{E}\big(\delta_x(X_T)U^i_T\hat{W}^j_T\big) = 0,$$
$$\mathbb{E}\big(\delta_x(X_T)\hat{W}^i_T\hat{W}^j_T\big) = \mathbb{E}\big(\delta_x(X_T)\big)T\delta_{ij} = Tp(x)\delta_{ij}.$$

Therefore

$$\mathbb{E}\Big\{\delta_x(X_T)Y_{ii',\theta}\Big\} = \mathbb{E}\Big\{\delta_x(X_T)U^i_TU^{i'}_T\Big\} + Tp(x)\mathbf{1}_{\{\theta=0\}}\delta_{ii'}$$

Therefore we finally obtain that

$$n^{\beta-\alpha(2+d)}\,\Big\|\sum_{i=1}^d \frac{\partial\phi_{h,x}}{\partial x_i}(X_T)(U^{n^\beta,i}_T + Z^i_{n^\beta,\theta})\Big\|_{L^2(\Omega)} \underset{n\to\infty}{\longrightarrow} \tilde{\sigma}^2,$$

from which the first assertion of the Lemma follows.

The second assertion is a consequence of the first. In fact using the tringular inequality, we have that

$$n^{\frac{\beta-\alpha(2+d)}{2}}\left|\Big\|\phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X^{n^\beta}_T + Z_{n^\beta,\theta})\Big\|_{L^2(\Omega)} - \Big\|\phi_{h,x}(X^{n^\beta}_T + Z_{n^\beta,\theta}) - \phi_{h,x}(X_T)\Big\|_{L^2(\Omega)}\right| \leq$$
$$n^{\frac{\beta-\alpha(2+d)}{2}}\Big\|\phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X_T)\Big\|_{L^2(\Omega)}.$$

As the first assertion is also valid for $\beta' \in (\beta, 1)$. We apply this first assertion noting that $\alpha \leq \frac{\beta}{d+2} < \frac{\beta'}{d+2}$ which gives

$$\lim_{n\to\infty} n^{\frac{\beta'-\alpha(2+d)}{2}}\|\phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X_T)\|_{L^2(\Omega)} = \tilde{\sigma}.$$

From here it follows that

$$n^{\frac{\beta-\alpha(2+d)}{2}}\Big\|\phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X_T)\Big\|_{L^2(\Omega)} \underset{n\to\infty}{\longrightarrow} 0.$$

From here the proof of the second assertion follows. □

Proof of Theorem 6.1. We have

$$n\big(V_n - p(x)\big) := \frac{1}{n^{\gamma_1-1}}\sum_{i=1}^{n^{\gamma_1}} \zeta^{n^\beta,h}_{T,i} + \frac{1}{n^{\gamma_2-1}}\sum_{i=1}^{n^{\gamma_2}} \tilde{\zeta}^{n,h}_{T,i} + n\big(\mathbb{E}\phi_{h,x}(X^n_T + Z_{n,\theta}) - p(x)\big)$$

with

$$\zeta^{n^\beta,h}_T = \phi_{h,x}(\hat{X}^{n^\beta}_T + \hat{Z}_{n^\beta,\theta}) - \mathbb{E}\,\phi_{h,x}(\hat{X}^{n^\beta}_T + \hat{Z}_{n^\beta,\theta})$$

and

$$\tilde{\zeta}^{n,h}_T = \phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X^{n^\beta}_T + Z_{n^\beta,\theta}) - \mathbb{E}\left\{\phi_{h,x}(X^n_T + Z_{n,\theta}) - \phi_{h,x}(X^{n^\beta}_T + Z_{n^\beta,\theta})\right\}.$$

From Theorem 4.1 and for $\gamma_1 = 2 + \alpha d$ we have that

$$\frac{1}{n^{\gamma_1 - 1}} \sum_{i=1}^{n^{\gamma_1}} \zeta_{T,i}^{n^\beta,h} \Rightarrow N(0,\sigma^2) \qquad \text{with} \quad \sigma^2 = \phi_2 p(x)$$

where $\phi_2 = \int_{\mathbb{R}^d} \phi^2(u)\,du$. Therefore to finish the proof it is enough to prove a central limit theorem for $\frac{1}{n^{\gamma_2-1}} \sum_{i=1}^{n^{\gamma_2}} \tilde{\zeta}_{T,i}^{n,h}$, as the random variables $\zeta_T^{n^\beta,h}$ are $\tilde{\zeta}_T^{n,h}$ independent. As in the proof of Theorem 4.1 we have that

$$\mathbb{E}\Big[\exp\big(\frac{iu}{n^{\gamma_2-1}} \sum_{k=1}^{n^{\gamma_2}} \tilde{\zeta}_{T,k}^{n,h}\big)\Big] = \Big[1 + \frac{1}{n^{\gamma_2}}\Big(\frac{-u^2}{2n^{\gamma_2-2}} \mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^2 + \mathbb{E}\,\tilde{C}_{n,h}(\omega)\Big)\Big]^{n^{\gamma_2}},$$

with

$$|\mathbb{E}\,\tilde{C}_{n,h}(\omega)| \le \frac{u^3}{6n^{2\gamma_2-3}} \mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^3.$$

Now we prove that

$$\frac{1}{n^{\gamma_2-2}} \mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^2 \underset{n\to\infty}{\longrightarrow} \tilde{\sigma}^2$$

and

$$\frac{1}{n^{2\gamma_2-3}} \mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^3 \underset{n\to\infty}{\longrightarrow} 0,$$

which will give as in the proof of Theorem 4.1

$$\frac{1}{n^{\gamma_2-1}} \sum_{i=1}^{n^{\gamma_2}} \tilde{\zeta}_{T,i}^{n,h} \to \tilde{\sigma} G$$

where G is a standard Gaussian random variable.

Let's start with the term $\mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^2$. We have that

$$\mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^2 = \mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right]^2 - \left\{\frac{C^s_{\phi,x}}{n} - \frac{C^s_{\phi,x}}{n^\beta} + o\Big(\frac{1}{n^\beta}\Big)\right\}^2,$$

where $C^s_{\phi,x}$ is the constant given in Theorem 3.1 associated to the kernel ϕ. Also from Lemma 6.1 and for $\gamma_2 = (d+2)\alpha + 2 - \beta$ we have that

$$\frac{1}{n^{\gamma_2-2}} \mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^2 \underset{n\to\infty}{\longrightarrow} \tilde{\sigma}^2.$$

On the other hand,

$$\mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^3 \le 4\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^3 + \left|\mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right]\right|^3$$

$$+ 3\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^2 \times \left|\mathbb{E}\left[\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right]\right|.$$

Also using Theorem 3.1 we obtain

$$\mathbb{E}\,|\tilde{\zeta}_T^{n,h}|^3 \le 4\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^3 + \left|\frac{C^s_{\phi,x}}{n} - \frac{C^s_{\phi,x}}{n^\beta} + o\Big(\frac{1}{n^\beta}\Big)\right|^3$$

$$+ 3\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^2 \left|\frac{C^s_{\phi,x}}{n} - \frac{C^s_{\phi,x}}{n^\beta} + o\Big(\frac{1}{n^\beta}\Big)\right|.$$

Applying again Lemma 6.1 we have that for $\gamma_2 = (d+2)\alpha + 2 - \beta$

$$\frac{1}{n^{2\gamma_2-3}}\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^2 \underset{n\to\infty}{\longrightarrow} 0.$$

Therefore it remains to prove that

$$\frac{1}{n^{2\gamma_2-3}}\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^3 \underset{n\to\infty}{\longrightarrow} 0. \tag{28}$$

As $\phi_{h,x}$ is a Lipschitz function with Lipschitz constant of c/h^{d+1} for $c>0$, we obtain

$$\begin{aligned}
\frac{1}{n^{2\gamma_2-3}}\mathbb{E}\left|\phi_{h,x}(X_T^n + Z_{n,\theta}) - \phi_{h,x}(X_T^{n^\beta} + Z_{n^\beta,\theta})\right|^3 &\le \frac{c}{n^{2\gamma_2-3}h^{3(d+1)}}\left[\mathbb{E}|X_T^n - X^{n^\beta}|^3 + \mathbb{E}|Z_{n,\theta} - Z_{n^\beta,\theta}|^3\right] \\
&\le \frac{c}{n^{2\gamma_2-3}h^{3(d+1)}} \times \frac{C_T}{n^{\frac{3\beta}{2}}} \\
&= \frac{c\,C_T}{n^{-(d-1)\alpha+1-\frac{\beta}{2}}} \to 0.
\end{aligned}$$

The last convergence is true if $0<\alpha<\beta/(d+2)$ and $0<\beta<2/3$. This finishes the proof of the Theorem. □

Like in the case of the Monte Carlo method one can interpret the previous result as follows: In order to approach the density $p(x)$ using a control variate method of the Romberg type with a global tolerance error of order $1/n$, the parameters needed to use the algorithm are $h = n^{-\alpha}$, $N_1 = n^{2+\alpha d}$ $N_2 = n^{(d+2)\alpha+2-\beta}$ with $\beta/(d+2) > \alpha \ge 1/s$ where s denotes the order of the superkernel ϕ. Therefore the complexity (number of calculations) needed for this algorithm is

$$\begin{aligned}
C_{RS} &= C \times mN_1 + (n+m)N_2 \\
&\simeq C \times n^{\beta+\alpha d+2} + n^{(d+2)\alpha-\beta+3}, \quad \text{where} \quad \beta/(d+2) > \alpha \ge 1/s.
\end{aligned}$$

For $\beta = \frac{1}{2} + \alpha$ we obtain that the complexity of the Romberg method is given by

$$C^\star_{RS} \simeq C \times n^{\frac{5}{2}+(d+1)\alpha}.$$

Here note that the optimal complexity for the Monte Carlo method is given by

$$C^\star_{MC} \simeq C \times n^{3+\alpha d}.$$

Therefore the Romberg control variate method reduces the complexity by a factor of of order $n^{1/2-\alpha}$. Therefore taking into account that $\beta/(d+2) > \alpha \ge 1/s$ we see that if one uses super-kernels of order $s > 2(d+1)$ we obtain a theoretical asymptotic optimal parameter choice of the method.

7 Appendix 1

In this appendix we prove some estimates that are useful to estimate the norms of the weights in the integration by parts formula. In order to simplify the notation we suppose that c is a positive constant being able to change from a line to another.

Lemma 7.1. *Under the above notations, we have that for all $k>1$, $p>1$ there exists positive constants k_2, p_1, p_2, γ_1 and γ_2 and a positive constant c independent of n, θ and F such that*

$$\|(\tilde\gamma^{-1}_{F+Z_{n,\theta}})^{ij}\|_{k,p} \le c\,\|(\det\tilde\gamma_{F+Z_{n,\theta}})^{-1}\|^{\gamma_1}_{p_1}\|F + Z_{n,\theta}\|^{\gamma_2}_{k_2,p_2}, \tag{29}$$

Proof. The proof is done by induction on k. The case $k=0$ is a direct consequence of the Cramer formula for the inverse of a given matrix.

In general, as $D^r\left\{\tilde{\gamma}_{F+Z_{n,\theta}}\tilde{\gamma}^{-1}_{F+Z_{n,\theta}}\right\}=0$ for any multi-index r, we have that

$$D^r(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{lm}=\sum_{i,j=1}^{d}\sum_{k\in A(r)-\{r\}}(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{li}D^{r-k}\tilde{\gamma}^{ij}_{F+Z_{n,\theta}}D^k(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{jm}.$$

Here $A(r)$ denotes all the subsets of indices of any order taken from elements of r. Then the result follows from the inductive hypothesis. □

Proposition 7.1. *Let $G\in\mathbb{D}^\infty(\tilde{W})$, then*

1. *Let $F\in(\mathbb{D}^\infty(W))^d$ such that $F+Z_{n,\theta}$ is a non-degenerate random vector. For $p>1$ and for all multi-index m we have*

$$\|\tilde{\mathbf{H}}_m(F,G)\|_p\le c\|G\|_{r,r'}\|(\det\tilde{\gamma}_{F+Z_{n,\theta}})^{-1}\|_a^{a'}\left[\|F\|_{b,l}^{b'}+\frac{1}{n^{(\frac{1}{2}+\theta)l'}}\right]$$

where c is a constant depending on p, m and d, whereas r, r', a, a', b, b', l and l' are parameters depending on m, p and d.

2. *Let $F_1,F_2\in(\mathbb{D}^\infty(W))^d$ such that $F_1+Z_{n,\theta}$ and $F_2+Z_{n,\theta}$ are non-degenerate random vectors. For a fixed multi-index m, any $k\ge1$ and $p>1$, there exists c, a positive constant depending on p, m and d, whereas k_i, s_i, β_i, p_i, γ_i, for $i=1,2$, k_0, s_0, γ_0, $\bar{k}_0$ and $\bar{s}_0$ are parameters depending on m, p and d such that*

$$\begin{aligned}\|\tilde{\mathbf{H}}_m(F_1,G)-\tilde{\mathbf{H}}_m(F_2,G)\|_{k,p}\le c\prod_{i=1}^{2}(1+\|F_i\|_{k_i,s_i}^{\gamma_i})(1+\|(\det\tilde{\gamma}_{F_i+Z_{n,\theta}})^{-1}\|_{p_i}^{\beta_i})&\\ \times\|F_1-F_2\|_{k_0,s_0}^{\gamma_0}\|G\|_{\bar{k}_0,\bar{s}_0}&\end{aligned}$$

3. *Let $F\in(\mathbb{D}^\infty(W))^d$ be a non-degenerate random vector. For a fixed multi-index m, any $k\ge1$ and $p>1$, there exists a constant c and parameters r_i, k_i, μ_i, for $i=1,2,3$ depending on p, m and d such that*

$$\|\tilde{\mathbf{H}}_m(F,G)-\mathbf{H}_m(F,G)\|_{k,p}\le\frac{c}{n^{(\frac{1}{2}+\theta)\mu}}\|G\|_{k_1,r_1}(1+\|F+Z_{n,\theta}\|_{k_2,r_2}^{\mu_2})(1+\|(\det\tilde{\gamma}_{F+Z_{n,\theta}})^{-1}\|_{r_3}^{\mu_3}).$$

Proof. Again the proof is done by induction on the length of the multi-index m. In fact, using the definition of $\tilde{\mathbf{H}}$ and the continuity of the adjoint operator δ, we have

$$\|\tilde{\mathbf{H}}_m(F,G)\|_{k,p}\quad\le\quad\sum_{r\in m}\sum_{j=1}^{d}\|\tilde{D}(F+Z_{n,\theta})^j\,\tilde{\mathbf{H}}_{m-\{r\}}(F,G)(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{rj}\|_{k+1,p}.$$

Then the proof finishes by using Hölder's inequality, Lemma 7.1 and the inductive hypothesis.

The proof of the second assertion is as the previous one, done by induction on the order of the multi-index m

$$\|\tilde{\mathbf{H}}_m(F_1,G)-\tilde{\mathbf{H}}_m(F_2,G)\|_{k,p}\leq\sum_{r\in m}\sum_{j=1}^{d}\|\tilde{D}(F_1-F_2)^j\,\tilde{\mathbf{H}}_{m-\{r\}}(F_1,G)(\tilde{\gamma}^{-1}_{F_1+Z_{n,\theta}})^{rj}\|_{k+1,p}$$

$$+\sum_{r\in m}\sum_{j=1}^{d}\|\tilde{D}(F_2+Z_{n,\theta})^j(\tilde{\mathbf{H}}_{m-\{r\}}(F_1,G)-\tilde{\mathbf{H}}_{m-\{r\}}(F_2,G))(\tilde{\gamma}^{-1}_{F_1+Z_{n,\theta}})^{rj}\|_{k+1,p}$$

$$+\sum_{r\in m}\sum_{j=1}^{d}\|\tilde{D}(F_2+Z_{n,\theta})^j\,\tilde{\mathbf{H}}_{m-\{r\}}(F_2,G)((\tilde{\gamma}^{-1}_{F_1+Z_{n,\theta}})^{rj}-(\tilde{\gamma}^{-1}_{F_2+Z_{n,\theta}})^{rj})\|_{k+1,p}.$$

For the first term one applies the Hölder's inequality, the first assertion and Lemma 7.1. For the second, Hölder's inequality, Lemma 7.1 and the inductive hypothesis. For the third, note that

$$(\tilde{\gamma}^{-1}_{F_1+Z_{n,\theta}})^{rj}-(\tilde{\gamma}^{-1}_{F_2+Z_{n,\theta}})^{rj}=\sum_{k,k'=1}^{d}(\tilde{\gamma}^{-1}_{F_2+Z_{n,\theta}})^{rk}\Big[(\tilde{\gamma}_{F_2+Z_{n,\theta}})^{kk'}-(\tilde{\gamma}_{F_1+Z_{n,\theta}})^{kk'}\Big](\tilde{\gamma}^{-1}_{F_1+Z_{n,\theta}})^{k'j}\quad(30)$$

Note that $\tilde{\gamma}_{F_1+Z_{n,\theta}}=\gamma_{F_1}+\bar{\gamma}_{Z_{n,\theta}}$ and $\tilde{\gamma}_{F_2+Z_{n,\theta}}=\gamma_{F_2}+\bar{\gamma}_{Z_{n,\theta}}$. Consequently, it follows that

$$(\tilde{\gamma}_{F_2+Z_{n,\theta}})^{kk'}-(\tilde{\gamma}_{F_1+Z_{n,\theta}})^{kk'}=\langle DF_2^k-DF_1^k,DF_2^{k'}\rangle_H+\langle DF_1^k,DF_2^{k'}-DF_1^{k'}\rangle_H.$$

From here the result follows.

In the same way as before, we prove the last relation for an index m. We have

$$\begin{aligned}\tilde{\mathbf{H}}_m(F,G)-\mathbf{H}_m(F,G)\;&=\;\sum_{j=1}^{m}\delta\Big(GDF^j\big[(\tilde{\gamma}^{-1}_{F+Z_{n,\theta}})^{mj}-(\tilde{\gamma}_F^{-1})^{mj}\big]\Big),\\&\quad+\frac{1}{n^{\frac{1}{2}+\theta}}\sum_{j=1}^{d}\bar{\delta}\Big(G(\gamma^{-1}_{F+Z_{n,\theta}})^{ij}\bar{D}\bar{W}^j_T\Big)\end{aligned}$$

Therefore the result follows applying (30) and the same arguments as in the previous proofs of assertions 1 and 2. □

8 Appendix 2

Proof of Lemma 3.2. We have that

$$\int_0^T\int_{\eta_n(u)}^u g_n(s,u)\,ds\,du=\int_0^T\int_{\eta_n(u)}^u g(s,u)\,ds\,du+\int_0^T\int_{\eta_n(u)}^u(g_n(s,u)-g(s,u))\,ds\,du.$$

In virtue of *ii)* we obtain that $R_n=o(1/n)$. Hence, we have

$$\begin{aligned}I_n:&=\int_0^T\int_{\eta_n(u)}^u g(s,u)\,ds\,du\\&=\int_0^T g(u,u)(u-\eta_n(u))\,du+\int_0^T du\int_{\eta_n(u)}^u(g(u,u)-g(s,u))\,ds.\end{aligned}$$

As g is uniformly continuous we have that for any $\varepsilon>0$ there exists $n_\varepsilon\in\mathbb{N}$ such that for all $n\geq n_\varepsilon$ we have

$$\sup_{0\leq u\leq T}\sup_{|s-u|\leq 1/n}|g(u,u)-g(s,u)|\leq\varepsilon.$$

Hence,

$$\Big|\int_0^T \int_{\eta_n(u)}^u (g(u,u) - g(s,u))\, ds\, du\Big| \le \frac{\varepsilon}{n}.$$

In addition, we have

$$\begin{aligned}
I'_n :&= \int_0^T g(u,u)(u - \eta_n(u))\, du \\
&= \int_0^T g(\eta_n(u), \eta_n(u))(u - \eta_n(u))\, du - \int_0^T \Big[g(\eta_n(u), \eta_n(u)) - g(u,u)\Big](u - \eta_n(u))\, du.
\end{aligned}$$

For the same reasons as before, for $n \ge n_\varepsilon$ we have

$$\Big|\int_0^T \Big[g(\eta_n(u), \eta_n(u)) - g(u,u)\Big](u - \eta_n(u))\, du\Big| \le \frac{\varepsilon}{n}.$$

Therefore, we have

$$\begin{aligned}
I''_n :&= \int_0^T g(\eta_n(u), \eta_n(u))(u - \eta_n(u))\, du \\
&= \frac{1}{2n^2} \sum_{l=0}^{n-1} g(l/n, l/n) \\
&= \frac{1}{2n} \int_0^T g(\eta_n(u), \eta_n(u))\, du \\
&= \frac{1}{2n} \int_0^T g(u,u)\, du + \frac{1}{2n} \int_0^T (g(\eta_n(u), \eta_n(u)) - g(u,u))\, du.
\end{aligned}$$

Similarly as before, for $n \ge n_\varepsilon$ we obtain

$$\frac{1}{n}\Big|\int_0^T (g(\eta_n(u), \eta_n(u)) - g(u,u))\, du\Big| \le \frac{\varepsilon}{n}.$$

We deduce that

$$\overline{\lim_n}\, n\Big|I_n - \frac{1}{2n}\int_0^T g(u,u)\, du\Big| \le \varepsilon.$$

Since $\varepsilon > 0$ is arbitrary, we conclude that

$$I_n = \frac{1}{2n}\int_0^T g(u,u)\, du + o(1/n).$$

□

Proof of Lemma 3.3. In order to prove the relation (12) is enough to prove that

$$\sup_{|u-s|\le\delta} |\Delta_{n,h}(u,s)| := \sup_{|u-s|\le\delta} |\Delta^1_{n,h}(u,s) + \Delta^2_{n,h}(s,u)| \to 0$$

with

$$\Delta^1_{n,h}(u,s) := \mathbb{E}\Big[\big(\psi_{h,x}\big(\zeta^n_\lambda + Z_{n,\theta}\big) - \mathbf{1}_{\{X_T > x\}}\big)\, \mathbf{H}_{m^+}\big(X_T, G^{r,j,k}_u\big)\Big]$$

and

$$\Delta^2_{n,h}(s,u) := \mathbb{E}\Big[\psi_{h,x}\big(\zeta^n_\lambda + Z_{n,\theta}\big)\big\{\tilde{\mathbf{H}}_{m^+}\big(\zeta^n_\lambda, G^{n,r,j,k}_{u,s}\big) - \mathbf{H}_{m^+}\big(X_T, G^{r,j,k}_u\big)\big\}\Big]$$

Since for every $p \ge 1$ we have that

$$\psi_{h,x}\big(\zeta^n_\lambda + Z_{n,\theta}\big) \xrightarrow{L^p} \mathbf{1}_{\{X_T > x\}}$$

and

$$\sup_{u\in[0,T]} \left\| \mathbf{H}_{m^+}\left(X_T, G_u^{r,j,k}\right) \right\|_p < \infty$$

then we deduce that

$$\sup_{|u-s|\leq\delta} |\Delta^1_{n,h}(u,s)| \to 0. \tag{31}$$

In addition we have

$$\sup_{|u-s|\leq\delta} |\Delta^2_{n,h}(s,u)| \leq c \sup_{|u-s|\leq\delta} \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{n,r,j,k}_{u,s}\right) - \mathbf{H}_{m^+}\left(X_T, G^{r,j,k}_u\right) \right\|_2$$

$$\begin{aligned}
\sup_{|u-s|\leq\delta\in[0,T]} |\Delta^2_{n,h}(s,u)| &\leq c \sup_{|u-s|\leq\delta} \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{n,r,j,k}_{u,s}\right) - \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{r,j,k}_u\right) \right\|_2 \\
&+ c \sup_{u\in[0,T]} \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{r,j,k}_u\right) - \tilde{\mathbf{H}}_{m^+}\left(X_T, G^{r,j,k}_u\right) \right\|_2 \\
&+ c \sup_{u\in[0,T]} \left\| \tilde{\mathbf{H}}_{m^+}\left(X_T, G^{r,j,k}_u\right) - \mathbf{H}_{m^+}\left(X_T, G^{r,j,k}_u\right) \right\|_2.
\end{aligned}$$

Note that

$$\left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{n,r,j,k}_{u,s}\right) - \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{r,j,k}_u\right) \right\|_2 = \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{n,r,j,k}_{u,s} - G^{r,j,k}_u\right) \right\|_2.$$

Since $\zeta^n_\lambda + Z_{n,\theta}$ is non-degenerate, we conclude using the second assertion of proposition 7.1 and properties (4) and (5), specific to the diffusion X and its Euler scheme X^n, that

$$\sup_{|u-s|\leq\delta} \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{n,r,j,k}_{u,s} - G^{r,j,k}_u\right) \right\|_2 \to 0, \quad (n\to\infty).$$

In the same way, since $\zeta^n_\lambda + Z_{n,\theta}$ and X_T are non-degenerate, we conclude using the second assertion of proposition 7.1 and relations (4) and (5), that

$$\sup_{u\in[0,T]} \left\| \tilde{\mathbf{H}}_{m^+}\left(\zeta^n_\lambda, G^{r,j,k}_u\right) - \tilde{\mathbf{H}}_{m^+}\left(X_T, G^{r,j,k}_u\right) \right\|_2 \to 0, \quad (n\to\infty).$$

Finally, according to the third assertion of proposition 7.1 we obtain that

$$\sup_{u\in[0,T]} \left\| \tilde{\mathbf{H}}_{m^+}\left(X_T, G^{r,j,k}_u\right) - \mathbf{H}_{m^+}\left(X_T, G^{r,j,k}_u\right) \right\|_2 \to 0, \quad (n\to\infty).$$

We conclude that

$$\sup_{|u-s|\leq\delta} |\Delta^2_{n,h}(u,s)| \to 0, \quad (n\to\infty).$$

□

Proof of Lemma 6.2. The case $d=1$ is trivial, so we will assume for the rest of the proof that $d\geq 2$. It is clear that the function $\partial_r Q$ is continous except at the origin. Since the random vector $\xi^{ii'jj'}_h \to O$ as $h\to 0$ a.s., the first assertion of the lemma follows. Now we prove the second assertion. We have that

$$\begin{aligned}
\mathbb{E}\left|\partial_r Q_d(X_T+\xi^{ii'jj'}_h - x)\right|^{1+\delta} &= E\Big\{\left|\partial_r Q_d(X_T+\xi^{ii'jj'}_h - x)\right|^{1+\delta} \mathbf{1}_{\{|X_T+\xi^{ii'jj'}_h - x|\leq 2|x|\}}\Big\} \\
&+ E\Big\{\left|\partial_r Q_d(X_T+\xi^{ii'jj'}_h - x)\right|^{1+\delta} \mathbf{1}_{\{|X_T+\xi^{ii'jj'}_h - x| > 2|x|\}}\Big\}.
\end{aligned} \tag{32}$$

• **Step 1:** For the first right term of the previous equality, we have

$$E\Big\{\big|\partial_r Q_d(X_T+\xi_h^{ii'jj'}-x)\big|^{1+\delta}\mathbf{1}_{\{|X_T+\xi_h^{ii'jj'}-x|\,\le\,2|x|\}}\Big\} = \int_{\mathbb{R}^d}\big|\partial_r Q_d(y-x)\big|^{1+\delta}\mathbf{1}_{\{|y-x|\,\le\,2|x|\}}p_h^{ii'jj'}(y)dy \quad (33)$$

where $p_h^{ii'jj'}$ denotes the density of the random vector $X_T+\xi_h^{ii'jj'}$. If we have that

$$\sup_h \sup_{|y-x|\le 2|x|} p_h^{ii'jj'}(y)\le C_x,$$

then it follows immediately that

$$\begin{aligned} E\Big\{\big|\partial_r Q_d(X_T+\xi_h^{ii'jj'}-x)\big|^{1+\delta}\mathbf{1}_{\{|X_T+\xi_h^{ii'jj'}-x|\,\le\,2|x|\}}\Big\} &\le C_x\int_{|y-x|\le 2|x|}\big|\partial_r Q_d(y-x)\big|^{1+\delta}dy \\ &= \int_{|y|\le 2|x|}\big|\partial_r Q_d(y)\big|^{1+\delta}dy. \end{aligned}$$

As $\big|\partial_r Q_d(y)\big|^{1+\delta}\le \frac{C_d}{|y|^{(d-1)(1+\delta)}}$. Therefore we obtain

$$\sup_h E\Big\{\big|\partial_r Q_d(X_T+\xi_h^{ii'jj'}-x)\big|^{1+\delta}\mathbf{1}_{\{|X_T+\xi_h^{ii'jj'}-x|\,\le\,2|x|\}}\Big\}<\infty, \quad (34)$$

for $\delta<(d-1)^{-1}$.

• **Step 2:** Now, we have to prove that

$$\sup_h \sup_{|y-x|\le 2|x|} p_h^{ii'jj'}(y)\le C_x.$$

We have that

$$p_h^{ii'jj'}(y)=\frac{1}{C_{ii'jj'}}\int_{\mathbb{R}^d}\varphi_{h,x}^{ii'jj'}(u-y)p(u)\,du,$$

where p denotes the density of X_T. Then it follows that

$$\begin{aligned} p_h^{ii'jj'}(y) &\le \sup_{u\in\mathbb{R}^d} p(u)\frac{1}{C_{ii'jj'}}\int_{\mathbb{R}^d}\varphi_{h,x}^{ii'jj'}(u-y)\,du \\ &= \sup_{u\in\mathbb{R}^d} p(u). \end{aligned}$$

Using the smoothness of p the result follows.

• **Step 3:** We denote by

$$I_+=\{1\le i\le d\,|\,y^i-x^i\ge 0\}$$

and

$$I_-=\{1\le i\le d\,|\,y^i-x^i< 0\}.$$

If $|y-x|\neq 0$, we have that for any $\gamma>0$ that

$$\begin{aligned} p(y) &= \mathbb{E}\Big[(-1)^{|I_-|}\prod_{i\in I_+}\mathbf{1}_{\{X_T^i-x^i>y^i-x^i\ge 0\}}\times\prod_{j\in I_-}\mathbf{1}_{\{X_T^j-x^j<y^j-x^j\le 0\}}\mathbf{H}_{(1,\dots,d)}(X_T,1)\Big] \\ &\le \Big[\mathbb{P}\big(|X_T-x|>|y-x|\big)\Big]^{\frac12}\Big\|\mathbf{H}_{(1,\dots,d)}(X_T,1)\Big\|_{L^2(\Omega)} \\ &\le \frac{\Big[\mathbb{E}\big|X_T-x\big|^{2\gamma}\Big]^{\frac12}}{|y-x|^{2\gamma}}\Big\|\mathbf{H}_{(1,\dots,d)}(X_T,1)\Big\|_{L^2(\Omega)} \\ &\le \frac{C_{x,d}}{|y-x|^{\gamma}}. \end{aligned}$$

• **Step 4:** Now we are able to deal with the second term of equality (32). We have

$$E\Big\{\big|\partial_r Q_d(X_T+\xi_h^{ii'jj'}-x)\big|^{1+\delta}\mathbf{1}_{\{|X_T+\xi_h^{ii'jj'}-x|\,>\,2|x|\}}\Big\}=$$
$$\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y+z-x)\big|^{1+\delta}\mathbf{1}_{\{|y+z-x|\,>\,2|x|\}}p(y)\varphi_{h,x}^{ii'jj'}(z)dz\,dy.$$

It follows that

$$\begin{aligned}
&E\Big\{\big|\partial_r Q_d(X_T+\xi_h^{ii'jj'}-x)\big|^{1+\delta}\mathbf{1}_{\{|X_T+\xi_h^{ii'jj'}-x|\,>\,2|x|\}}\Big\}\\
&=\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}p(y-z+x)\varphi_{h,x}^{ii'jj'}(z)dz\,dy\\
&=\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}p(y-\theta h+2x)\varphi^{ii'jj'}(\theta)d\theta\,dy
\end{aligned}$$

• **First case:**
If $|y-\theta h+x|>|y+x|/2$ then using the step 3 result's we get

$$p(y-\theta h+2x)\le\frac{C_{x,d}}{|y-\theta h+x|^{\gamma}}\le\frac{2C_{x,d}}{|y+x|^{\gamma}}.$$

Therefore

$$\begin{aligned}
&\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}\mathbf{1}_{\{|y-\theta h+x|\,>\,|y+x|/2\}}p(y-\theta h+2x)\varphi^{ii'jj'}(\theta)d\theta\,dy\\
&\le 2C_{x,d}\int_{|y|>2|x|}\frac{\big|\partial_r Q_d(y)\big|^{1+\delta}}{|y+x|^{\gamma}}\,dy\\
&\le 2C_{x,d}\int_{|y|>2|x|}\frac{1}{\Big|1-\frac{|x|}{|y|}\Big|^{\gamma}}\times\frac{\big|\partial_r Q_d(y)\big|^{1+\delta}}{|y|^{\gamma}}\,dy\\
&\le 2^{\gamma+1}C_{x,d}\int_{|y|>2|x|}\frac{\big|\partial_r Q_d(y)\big|^{1+\delta}}{|y|^{\gamma}}\,dy<\infty\quad\text{since}\quad\delta>\frac{1-\gamma}{d-1}.
\end{aligned}$$

• **Second case:**
If $|y-\theta h+x|\le|y+x|/2$, then we have

$$\begin{aligned}
&\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}\mathbf{1}_{\{|y-\theta h+x|\,\le\,|y+x|/2\}}p(y-\theta h+2x)\varphi^{ii'jj'}(\theta)d\theta\,dy\\
&\le\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}\mathbf{1}_{\{|y+x|/2h\,\le\,|\theta|\,\le\,3|y+x|/2h\}}p(y-\theta h+2x)\varphi^{ii'jj'}(\theta)d\theta\,dy.
\end{aligned}$$

Using the assumption $\varphi^{ii'jj'}(\theta)\le c/|\theta|^{\gamma}$, for a given constant $c>0$ together with the relation (34), we obtain

$$\begin{aligned}
&\int_{\mathbb{R}^d\times\mathbb{R}^d}\big|\partial_r Q_d(y)\big|^{1+\delta}\mathbf{1}_{\{|y|\,>\,2|x|\}}\mathbf{1}_{\{|y-\theta h+x|\,\le\,|y+x|/2\}}p(y-\theta h+2x)\varphi^{ii'jj'}(\theta)d\theta\,dy\\
&\le c\times C_d\int_{\mathbb{R}^d\times\mathbb{R}^d}\Big[\frac{1}{|y|}\Big]^{(1+\delta)(d-1)}\mathbf{1}_{\{|\theta|\,\ge\,|x|/2h,\ |y|\,<\,2|x|\}}\frac{p(y-\theta h+2x)}{|\theta|^{\gamma}}d\theta\,dy.\\
&\le C'_{x,d}\int_{\mathbb{R}^d\times\mathbb{R}^d}\mathbf{1}_{\{|\theta|\,\ge\,|x|/2h\}}\frac{p(y-\theta h+2x)}{|\theta|^{\gamma}}d\theta\,dy\\
&=C'_{x,d}\int_{|\theta|\ge|x|/2h}\frac{1}{|\theta|^{\gamma}}d\theta\int_{\mathbb{R}^d}p(u)\,du<\infty,\ \text{where } C'_{x,d}\text{ is a positive constant.}
\end{aligned}$$

Which completes the lemma proof. □

9 Appendix 3

In the following we prove lemmas 5.1, 5.2 and 5.3.

Proof of Lemma 5.1. The proof uses the same ideas of Jacod and Protter (1998). Note that for $0 \le t < t' \le T$, the sequence $\left(\sqrt{n}\int_t^{t'}(W_s^j - W_{\eta_n(s)}^j)\,ds\right)_{n\in\mathbb{N}}$ tends to 0 in $L^2(\Omega)$. In fact, we have

$$\mathbb{E}\Big(\int_t^{t'}(W_s^j - W_{\eta_n(s)}^j)\,ds\Big)^2 \le \frac{c}{n^2}, \quad c>0.$$

Therefore we have

$$\int_0^T H_s^{i,n}(W_s^j - W_{\eta_n(s)}^j)\,ds = \int_0^T (H_s^{i,n} - H_{m,s}^{i,n})(W_s^j - W_{\eta_n(s)}^j)\,ds + \int_0^T H_{m,s}^{i,n}(W_s^j - W_{\eta_n(s)}^j)\,ds$$

with

$$H_m^{i,n} = \sum_{k=1}^m H_{\frac{kt}{m}}^{i,n}\mathbf{1}_{]\frac{(k-1)T}{m},\frac{kT}{m}]}.$$

It follows that

$$\begin{aligned}\Big|\int_0^T H_s^{i,n}(W_s^j - W_{\eta_n(s)}^j)\,ds\Big| \;&\le\; \sup_{0<s\le T}|H_s^{i,n} - H_{m,s}^{i,n}|\int_0^T |W_s^j - W_{\eta_n(s)}^j| \\ &\quad + \Big|\int_0^T H_{m,s}^{i,n}(W_s^j - W_{\eta_n(s)}^j)\,ds\Big|.\end{aligned}$$

Since the sequence $\sqrt{n}\int_0^T |W_s^j - W_{\eta_n(s)}^j|\,ds$ is tight, we deduce easily the lemma. □

Proof of Lemma 5.2. We denote by

$$H\diamond K^n = \int_0^T H_s\Big(\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}}K_u^n\,dW_u^j\Big)ds$$

and suppose in a first time that H is deterministic then

$$\begin{aligned}H\diamond K^n &= \int_0^T K_u^n\Big(\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}}H_s\,ds\Big)dW_u^j \\ &= \int_0^T K_u^n\Big(\int_{\eta_n(u)}^u H_s\,ds\Big)dW_u^j.\end{aligned}$$

It follows that

$$\begin{aligned}\|H\diamond K^n\|_{L^2(\Omega)}^2 &= \mathbb{E}\int_0^T (K_u^n)^2\Big(\int_{\eta_n(u)}^u H_s\,ds\Big)^2 du \\ &\le |H|_\infty^2\mathbb{E}\int_0^T (K_u^n)^2(u-\eta_n(u))^2\,du \\ &\le \frac{|H|_\infty^2 T^2}{n^2}\mathbb{E}\int_0^T (K_u^n)^2\,du,\end{aligned}$$

and consequently $(\sqrt{n}H\diamond K^n)_{n\in\mathbb{N}}$ tends to 0 in $L^2(\Omega)$. Now let H to be arbitrary. We have that $H\in\mathscr{C}([0,T])$, so there exists a sequence $H^l\in\mathscr{C}([0,T])$ of piecewise functions such that $|H-H^l|_\infty\to 0$ and $|H^l|_\infty \le l$ *a.s.*. We have

$$|(H-H^l)\diamond K^n| \le |H-H^l|_\infty\int_0^T\Big|\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}}K_u^n\,dW_u^j\Big|ds.$$

It is obvious that the sequence

$$\Big(\sqrt{n}\int_0^T\Big|\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}}K^n_u\,dW^j_u\Big|ds\Big)_{n\in\mathbb{N}}$$

is tight, (because it is bounded in L^2). Consequently:

$$\begin{aligned}\mathbb{P}\Big(\sqrt{n}|H\diamond K^n|\ge\varepsilon\Big) &\le \mathbb{P}\Big(\sqrt{n}|(H-H^l)\diamond K^n|\ge\frac{\varepsilon}{2}\Big)+\mathbb{P}\Big(\sqrt{n}|H^l\diamond K^n|\ge\frac{\varepsilon}{2}\Big)\\ &\le \mathbb{P}\Big(|H-H^l|_\infty\ge\frac{\varepsilon}{2\delta}\Big)\\ &\quad+\mathbb{P}\Big(\sqrt{n}\int_0^T\Big|\int_0^T \mathbf{1}_{\{\eta_n(u)\le s\le u\}}K^n_u\,dW^j_u\Big|ds\ge\delta\Big)\\ &\quad+\mathbb{P}\Big(\sqrt{n}|H^l\diamond K^n|\ge\frac{\varepsilon}{2}\Big).\end{aligned}$$

For a fixed l and for a good choice of δ and n we obtain that for a given $\eta>0$,

$$\limsup_{n\to\infty}\mathbb{P}\Big(\sqrt{n}|H\diamond K^n|\ge\varepsilon\Big)\le\eta+\mathbb{P}\Big(|H-H^l|_\infty\ge\frac{\varepsilon}{2\delta}\Big).$$

Since η is arbitrary and $|H-H^l|_\infty\to0$ *a.s.*, we conclude that

$$\limsup_{n\to\infty}\mathbb{P}\Big(\sqrt{n}|H\diamond K^n|\ge\varepsilon\Big)=0.$$

Which completes the proof. □

Proof of Lemma 5.3. We split the proof of the lemma into two steps
• **Step 1:** We suppose first H^i, K^i and L^i are deterministic. Then we have:

$$\begin{aligned}\int_0^T K^i_s\Big(\sum_{j=1}^q\int_s^T \xi^{ij}_{s,u}\bar{U}^n_u\,dW^j_u\Big)ds &= \int_0^T\Big(\int_0^u K^i_s\xi^{ij}_{s,u}ds\Big)\bar{U}^n_u dW^j_u\\ &=\sum_{j=1}^q\int_0^T \bar{K}^{ij}_u\bar{U}^n_u\,dW^j_u\end{aligned}$$

with $\bar{K}^{ij}_u=\int_0^u K^i_s\xi^{ij}_{s,u}\,ds$. In the same manner:

$$\int_0^T L^i_s\Big(\int_s^T\sum_{j,k=1}^q \zeta^{ijk}_{s,u}\,d\check{W}^{n,kj}_u\Big)ds=\sum_{j,k=1}^q\int_0^T \bar{L}^{ijk}_u\,d\check{W}^{n,kj}_u,$$

with $\bar{L}^{ijk}_u=\int_0^u L^i_s\zeta^{ijk}_{s,u}ds$. It remains to prove that

$$\Big(\bar{U}^n_T,\int_0^T H^i_s\bar{U}^n_s\,ds,\sum_{j=1}^q\int_0^T\bar{K}^{ij}_u\bar{U}^n_u\,dW^j_u,\sum_{j,k=1}^q\int_0^T\bar{L}^{ijk}_u\,d\check{W}^{n,kj}_u\Big)$$

stably converge in law to

$$\Big(U_T,\int_0^T H^i_sU_s\,ds,\sum_{j=1}^q\int_0^T\bar{K}^{ij}_uU_u\,dW^j_u,\sum_{j,k=1}^q\int_0^T\bar{L}^{ijk}_u\,d\check{W}^{kj}_u\Big)$$

Since the process H^i is deterministic and the processes $\bar{K}^{ij}$ and $\bar{L}^{ijk}$ are continuous adapted we deduce, using an approximation argument, that proving the convergence above can be carried into proving that $\sum_{i=1}^m Z_iV^n_i$ stably converge in law to $\sum_{i=1}^m Z_iV_i$ where $Z_1,\dots,Z_m$ are random matrices and $(V^n_1,\dots,V^n_m)$ are random vectors converging stably to $(V_1,\dots,V_m)$. This is a classical property of the stable convergence. In fact, $(Z, Z_1, \ldots, Z_u, V_1^n, \ldots, V_m^n)$ converge to $(Z, Z_1, \ldots, Z_u, V_1, \ldots, V_m)$, it follows that $(Z, \sum_{i=1}^m Z_i V_i^n)$ stably converge in law to $(Z, \sum_{i=1}^m Z_i V_i)$, (see Jacod and Shiryaev (2003) chapter VIII §5.c and Theorems 2.3 and 3.2 in Jacod and Protter (1998)).

• **Step 2:** Suppose now H^i, K^i and L^i are arbitrary. Since the processes H^i, K^i and L^i have continuous trajectories on $[0,T]$, we can approach them by three piecewise functions H_l^i, K_l^i, L_l^i. In the following we introduce the following notations

$$H^i.\bar{U}^n = \int_0^T H_s^i \bar{U}_s^n \, ds \qquad K^i \star U^n = \int_0^T K_s^i \Big(\int_s^T \sum_{j=1}^q \xi_{s,u} \bar{U}_u^n dW_u^j \Big) ds$$

$$(L^i | \check{W}^{n,kj}) = \int_0^T L_s^i \Big(\int_s^T \sum_{j,k=1}^n \zeta_{s,u}^{ijk} \, d\check{W}_u^{n,kj} \Big) ds.$$

We have $\|H^i.U^n - H_l^i.U^n\| \le |H^i - H_l^i|_\infty \int_0^T \|\bar{U}_s^n\| ds$ where $|\ |_\infty$ denotes the uniform norm on the space $\mathscr{C}([0,T])$. Similarly, we have

$$\|K^i \star U^n - K_l^i \star U^n\| \le |K^i - K_l^i|_\infty \int_0^T \Big\| \sum_{j=1}^q \int_s^T \xi_{s,u}^{ij} \bar{U}_u^n \, dW_u^j \Big\| ds$$

and

$$\|(L^i | \check{W}^{n,kj}) - (L_l^i | \check{W}^{n,kj})\| \le |L^i - L_l^i|_\infty \int_0^T \Big\| \int_s^T \sum_{j,k=1}^q \zeta_{s,u}^{ijk} \, d\check{W}_u^{n,kj} \Big\| ds.$$

Consequently, in order to prove the statement of the lemma, we have just to prove the tightness of

$$\int_0^T \|\bar{U}_s^n\| \, ds, \qquad P_n^i = \int_0^T \Big\| \int_s^T \sum_{j=1}^q \xi_{s,u}^{ij} \bar{U}_u^n \, dW_u^j \Big\| \, ds$$

$$\text{and} \quad Q_n^i = \int_0^T \Big\| \int_s^T \sum_{j,k=1}^q \zeta_{s,u}^{ijk} \, d\check{W}_u^{n,kj} \Big\| ds.$$

The tightness of the sequence $\int_0^T \|\bar{U}_s^n\| \, ds$, follows from the convergence of the law of $\bar{U}^n$. For P_n^i and Q_n^i, this is a consequence of the hypothesis on $\xi_{s,u}^{ij}$, $\zeta_{s,u}^{ijk}$. In fact :

$$\begin{aligned}
\|P_n^i\|_{L^2(\Omega)} &\le \int_0^T \Big\| \int_s^T \sum_{j=1}^q \xi_{s,u}^{ij} \bar{U}_u^n \, dW_u^j \Big\|_{L^2(\Omega)} ds \\
&= \int_0^T \Big\| \Big(\int_s^T \| \sum_{j=1}^q \xi_{s,u}^{ij} \bar{U}_u^n \|^2 du \Big)^{1/2} \Big\|_{L^2(\Omega)} ds \\
&\le \sqrt{T} \Big(\mathbb{E} \int_0^T ds \int_s^T du \| \sum_{j=1}^q \xi_{s,u}^{ij} \bar{U}_u^n \|^2 \Big)^{1/2} \\
&\le q\sqrt{T} \Big(\mathbb{E} \int_0^T du \Big[\int_0^u ds \max_j \|\xi_{s,u}^{ij}\|^2 \Big] \|\bar{U}_u^n\|^2 \Big)^{1/2},
\end{aligned}$$

Using that $\sup_n \mathbb{E} \int_0^T |\bar{U}_u^n|^q du < \infty$ for $q \ge 1$ and that

$$\mathbb{E} \int_0^T du \int_0^u ds \Big(\max_j \|\xi_{s,u}^{ij}\|^p \Big) < \infty \qquad \text{for} \quad p > 2.$$

we obtain that

$$\sup_n \|P_n\|_2 < \infty.$$

In the same manner we obtain that $\sup_n \|Q_n\|_{L^2(\Omega)} < \infty$ which completes the proof of the lemma. □

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